

# Models of Classical Test Theory

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April 3, 2009

Model of Essentially  $\tau$ -Equivalent Tests and  
Model of  $\tau$ -Congeneric Tests

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**Model of Essentially  $\tau$ -Equivalent Tests — Assumptions**

Definition: Assumptions (a<sub>2</sub>) and (b)

- (a<sub>2</sub>) essential  $\tau$ -equivalence:  $\tau_i = \lambda_{ij} + \tau_j$
- (b) uncorrelated errors  $Cov(\varepsilon_i, \varepsilon_j) = 0, \quad i \neq j$

First implication:

(a<sub>2</sub>) implies the existence of a latent variable  $\eta$  such that:

$$\tau_i = \lambda_i + \eta \quad \text{and} \quad Y_i = \lambda_i + \eta + \varepsilon_i \tag{1}$$

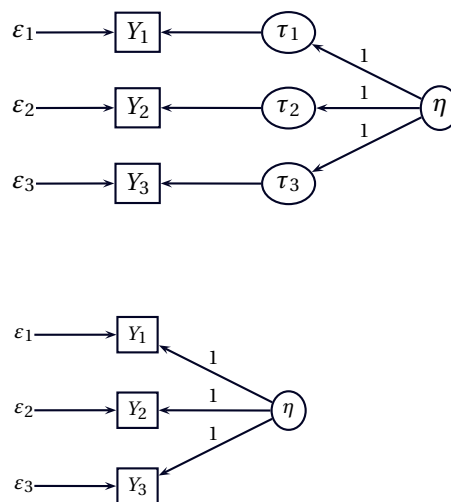
for all  $i = 1, \dots, m$

*Proof:*

Defining  $\eta := \tau_1$  and  $\lambda_i := \lambda_{i1}$ , Equation (1) follows from (a<sub>2</sub>).

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**Path diagram**



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**Implied Covariance Structure**

$$\begin{aligned} Cov(Y_i, Y_j) &= Cov(\lambda_1 + \eta + \varepsilon_1, \lambda_2 + \eta + \varepsilon_2) \\ &= Cov(\eta, \eta) + Cov(\eta, \varepsilon_2) + Cov(\varepsilon_1, \eta) + Cov(\varepsilon_1, \varepsilon_2) \\ &= Var(\eta) \\ &= \sigma_\eta^2 \end{aligned}$$

$$\begin{aligned} Var(Y_i) &= Var(\eta) + Var(\varepsilon_i) \\ &= \sigma_\eta^2 + \sigma_{\varepsilon_i}^2 \end{aligned}$$

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## Implied Covariance Matrix

Implied covariance matrix for 3 essentially  $\tau$ -equivalent tests:

$$\begin{bmatrix} \sigma_{\eta}^2 + \sigma_{\varepsilon_1}^2 & \sigma_{\eta}^2 & \sigma_{\eta}^2 \\ \sigma_{\eta}^2 & \sigma_{\eta}^2 + \sigma_{\varepsilon_2}^2 & \sigma_{\eta}^2 \\ \sigma_{\eta}^2 & \sigma_{\eta}^2 & \sigma_{\eta}^2 + \sigma_{\varepsilon_3}^2 \end{bmatrix}$$

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## Testability

Testability in the total population:

$$\text{Cov}(Y_i, Y_j) = \sigma_{\eta}^2, \quad i \neq j$$

Testability in each subpopulation  $s$ :

$$\text{Cov}^{(s)}(Y_i, Y_j) = \sigma_{\eta}^{(s)2}, \quad i \neq j$$

$$\begin{aligned} E^{(s)}(Y_i) - E^{(s)}(Y_j) &= E^{(s)}(Y_i - Y_j) \\ &= \lambda_{ij} \end{aligned}$$

The constant  $\lambda_{ij}$  is the same in all subpopulations.

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## Fixing Scales and Identification

Identification

$$\text{Var}(\eta) = \text{Cov}(Y_i, Y_j), \quad i \neq j$$

$$\text{Var}(\varepsilon_i) = \text{Var}(Y_i) - \text{Cov}(Y_i, Y_j), \quad i \neq j$$

$$\text{Rel}(Y_i) = \frac{\text{Cov}(Y_i, Y_j)}{\text{Var}(Y_i)}, \quad i \neq j$$

Fixing the scale of  $\eta$

First possibility:

$$\text{Fixing } E(\eta) = 0 \quad \text{implies} \quad \lambda_i = E(Y_i)$$

Second possibility:

$$\text{Fixing } \lambda_1 = 0 \quad \text{implies} \quad E(\eta) = E(Y_1) \quad \text{and} \quad \lambda_i = E(Y_i) - E(Y_1)$$

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### Uniqueness of Theoretical Concepts

$\eta$  and the coefficients  $\lambda_i$  are defined by the assumptions uniquely *up to translations*.

This means: the latent variable  $\eta$  and the coefficients  $\lambda_i$  are measured on a *difference scale*.

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### Meaningfulness

Under translations of  $\eta$  and the coefficients  $\lambda_i$ , the following propositions have invariant truth values ('true' or 'false'):

- $\eta(\omega_1) - \eta(\omega_2)$ , for all  $\omega_1, \omega_2 \in \Omega$
- $\lambda_i - \lambda_j$
- $Var(\eta)$
- $Rel(Y_i)$

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### Test Lengthening

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#### Spearman-Brown Formula and Cronbach's $\alpha$

**Test Lengthening**  $S = Y_1 + \dots + Y_m$

Assumptions (a<sub>2</sub>), (b) and (c) imply:

Spearman-Brown Formula:  $Rel(S) = \frac{m \cdot Rel(Y)}{1 + (m - 1) \cdot Rel(Y)}$

Assumption (b) (uncorrelated errors) implies that  $\alpha$  is a *lower bound for the reliability* of  $S = Y_1 + \dots + Y_m$ :

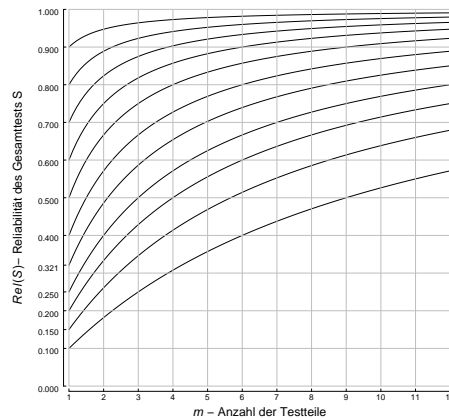
$$Cronbach's \alpha := \frac{m}{m-1} \cdot \left[ 1 - \frac{\sum_{i=1}^m Var(Y_i)}{Var(S)} \right] \quad (2)$$

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## Reliability

Reliability of the sum scores  $S := \sum_{i=1}^m Y_i$  as a function of test lengthening by factor  $m$  and of the reliability of the original variables  $Y_i$ , under the assumptions ( $a_1$ ) [or ( $a_2$ )], (b) and (c).



Grafische Darstellung der Spearman-Brown-Formel

## Model of $\tau$ -Congeneric Tests

### Model of $\tau$ -Congeneric Tests — Assumptions

Assumptions:

- ( $a_3$ )  $\tau$ -congenericity  
$$\tau_i = \lambda_{i0} + \lambda_{i1} \tau_j, \quad \lambda_{i0}, \lambda_{i1} \in \mathbb{R}, \quad \lambda_{i1} > 0$$
- (b) uncorrelated errors:  $Cov(\varepsilon_i, \varepsilon_j) = 0, \quad i \neq j$

( $a_3$ ) implies the existence of a latent variable  $\eta$  and coefficients  $\lambda_{i0}$  and  $\lambda_{i1}$  such that:

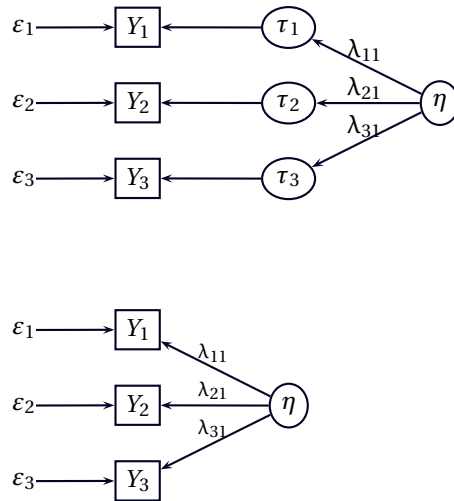
$$\tau_i = \lambda_{i0} + \lambda_{i1} \eta \quad \text{and, therefore,} \quad Y_i = \lambda_{i0} + \lambda_{i1} \eta + \varepsilon_i$$

*Proof:*

Just fix  $j = 1$  in ( $a_3$ ) and define:

$$\eta := \tau_1, \quad \lambda_{i0} := \lambda_{i10} \quad \text{and} \quad \lambda_{i1} := \lambda_{i11}$$

### Model of $\tau$ -Congeneric Tests — Path Diagram



### Illustration

**Tabelle.** Illustration of the relationship between the manifest and latent variables in the model of  $\tau$ -congeneric tests

Person	True scores		Latent scores	Observed scores		error scores		$P(Y_i = y_i   U = u)$
	$\tau_1$	$\tau_2$	$\eta$	$Y_1$	$Y_2$	$\epsilon_1$	$\epsilon_2$	
1	12	23	34	10	20	-2	-3	1/3
				12	24	0	1	1/3
				14	25	2	2	1/3
2	10	20	30	7	15	-3	-5	1/3
				9	22	-1	2	1/3
				14	23	4	3	1/3
3	8	17	26	3	14	-5	-3	1/3
				10	15	2	-2	1/3
				11	22	3	5	1/3

*Anmerkung:* Fictive numbers. Each of the three Persons has its own (intra-individual) distribution of the  $Y$ -variables, but only a single score on each of the latent variables  $\tau_i$  and  $\eta$ .

### Uniqueness

( $a_3$ ) implies the existence of a latent variable  $\eta$  and coefficients  $\lambda_{ij0}$  and  $\lambda_{ij1}$  such that:

$$(1) \quad \tau_i = \lambda_{i0} + \lambda_{i1} \eta \quad \text{and} \quad (2) \quad Y_i = \lambda_{i0} + \lambda_{i1} \eta + \varepsilon_i$$

The latent variable  $\eta$  and the coefficients  $\lambda_{i0}$  and  $\lambda_{i1}$  are uniquely defined only up to *linear transformations*. Hence, the equations above also hold for each variable  $\eta' := \alpha + \beta \eta$ , although with new coefficients  $\lambda'_{i0} = \lambda_{i0} - \frac{\lambda_{i1} \alpha}{\beta}$  and

$$\lambda'_{i1} = \frac{\lambda_{i1}}{\beta} .$$

*Proof:*

Inserting  $\eta' := \alpha + \beta \eta$  as well as  $\lambda'_{i0}$  and  $\lambda'_{i1}$  into Equation (1) yields

$$\begin{aligned} \tau_i &= \lambda'_{i0} + \lambda'_{i1} \eta' \\ &= \left( \lambda_{i0} - \frac{\lambda_{i1}}{\beta} \alpha \right) + \frac{\lambda_{i1}}{\beta} (\alpha + \beta \eta) \\ &= \lambda_{i0} + \lambda_{i1} \eta \end{aligned}$$

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### Fixing the Scale

In order to have a uniquely defined scale, we have to fix location and units of measurement, for instance:

$$\text{either by: } E(\eta) = 0 \quad \text{and} \quad \text{Var}(\eta) = 1$$

$$\text{or by: } \lambda_{i0} = 0 \quad \text{and} \quad \lambda_{i1} = 1$$

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### Implied Covariance Structure

$$\begin{aligned} \text{Cov}(Y_i, Y_j) &= \text{Cov}(\lambda_{i0} + \lambda_{i1} \eta + \varepsilon_i, \lambda_{j0} + \lambda_{j1} \eta + \varepsilon_j) \\ &= \lambda_{i1} \lambda_{j1} \text{Var}(\eta) \quad \text{for } i \neq j \end{aligned}$$

$$\begin{aligned} \text{Var}(Y_i) &= \text{Var}(\lambda_{i0} + \lambda_{i1} \eta + \varepsilon_i) \\ &= \lambda_{i1}^2 \text{Var}(\eta) + \text{Var}(\varepsilon_i) \end{aligned}$$

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### Implied Covariance Matrix for Three $\tau$ -Congeneric Tests

$$\begin{bmatrix} \lambda_{11}^2 \sigma_\eta^2 + \sigma_{\varepsilon_1}^2 & & \\ \lambda_{11} \lambda_{21} \sigma_\eta^2 & \lambda_{21}^2 \sigma_\eta^2 + \sigma_{\varepsilon_2}^2 & \\ \lambda_{11} \lambda_{31} \sigma_\eta^2 & \lambda_{21} \lambda_{31} \sigma_\eta^2 & \lambda_{31}^2 \sigma_\eta^2 + \sigma_{\varepsilon_3}^2 \end{bmatrix}$$

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### Identification of Loadings

$$\frac{\text{Cov}(Y_i, Y_j) \cdot \text{Cov}(Y_i, Y_k)}{\text{Cov}(Y_j, Y_k)} = \frac{\lambda_{i1} \lambda_{j1} \text{Var}(\eta) \cdot \lambda_{i1} \lambda_{k1} \text{Var}(\eta)}{\lambda_{j1} \lambda_{k1} \text{Var}(\eta)}$$
$$= \lambda_{i1}^2 \text{Var}(\eta) \quad i \neq j, \quad i \neq k, \quad j \neq k$$

If we fix  $\text{Var}(\eta) = 1$ , then:

$$\lambda_{i1} = \sqrt{\frac{\text{Cov}(Y_i, Y_j) \cdot \text{Cov}(Y_i, Y_k)}{\text{Cov}(Y_j, Y_k)}} \quad i \neq j, \quad i \neq k, \quad j \neq k$$

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### Identification of Other Parameters

If we fix the scale of  $\eta$  by

$$E(\eta) = 0 \quad \text{and} \quad \text{Var}(\eta) = 1,$$

then:

$$\text{Var}(\varepsilon_i) = \text{Var}(Y_i) - \lambda_{i1}^2$$

$$\text{Rel}(Y_i) = \frac{\lambda_{i1}^2}{\text{Var}(Y_i)}$$

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### Testability in the total Population

We can test the model of  $\tau$ -congeneric tests in the total population by testing the hypothesis:

$$\frac{\text{Cov}(Y_i, Y_k)}{\text{Cov}(Y_j, Y_k)} = \frac{\text{Cov}(Y_i, Y_l)}{\text{Cov}(Y_j, Y_l)}$$

This can be seen by inserting the implied covariance structure for each of the four covariances:

$$\frac{\lambda_{i1} \lambda_{k1} \text{Var}(\eta)}{\lambda_{j1} \lambda_{k1} \text{Var}(\eta)} = \frac{\lambda_{i1} \lambda_{l1} \text{Var}(\eta)}{\lambda_{j1} \lambda_{l1} \text{Var}(\eta)} \quad i \neq k, \quad i \neq l, \quad j \neq k, \quad j \neq l$$

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## Testability Using the Implied Mean Structure

The model of  $\tau$ -congeneric tests implies:

$$\frac{E^{(1)}(Y_i) - E^{(2)}(Y_i)}{E^{(1)}(Y_j) - E^{(2)}(Y_j)} = \frac{E^{(3)}(Y_i) - E^{(4)}(Y_i)}{E^{(3)}(Y_j) - E^{(4)}(Y_j)}$$

This can be seen as follows: Inserting the model equations into  $E^{(s)}(Y_i)$  yields:

$$E^{(s)}(\lambda_{i0} + \lambda_{i1}\eta + \varepsilon_i) = \lambda_{i0} + \lambda_{i1}E^{(s)}(\eta) \quad \text{for each subpopulation } s = 1, \dots, 4,$$

Inserting this result into the first equation yields:

$$\frac{\lambda_{i0} + \lambda_{i1}E^{(1)}(\eta) - [\lambda_{i0} + \lambda_{i1}E^{(2)}(\eta)]}{\lambda_{j0} + \lambda_{j1}E^{(1)}(\eta) - [\lambda_{j0} + \lambda_{j1}E^{(2)}(\eta)]} = \frac{\lambda_{i0} + \lambda_{i1}E^{(3)}(\eta) - [\lambda_{i0} + \lambda_{i1}E^{(4)}(\eta)]}{\lambda_{j0} + \lambda_{j1}E^{(3)}(\eta) - [\lambda_{j0} + \lambda_{j1}E^{(4)}(\eta)]}$$
$$\frac{\lambda_{i1}}{\lambda_{j1}} = \frac{\lambda_{i1}}{\lambda_{j1}}$$

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## Meaningfulness

For instance, propositions about

- $\frac{\eta(\omega_1) - \eta(\omega_2)}{\eta(\omega_3) - \eta(\omega_4)}$
- $\frac{\lambda_{i1}}{\lambda_{j1}}$
- $\lambda_{i1}^2 \text{Var}(\eta)$

have *invariant truth values* under the admissible transformations (here: linear transformations). The first equation means that  $\eta$  is measured on an *interval scale*.

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