

General Latent Variable Modeling Using Mplus Version 3

Block 1: Structural Equation Modeling

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Mplus: www.statmodel.com

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Program Background

- Inefficient dissemination of statistical methods:
 - Many good methods contributions from biostatistics, psychometrics, etc are underutilized in practice
- Fragmented presentation of methods:
 - Technical descriptions in many different journals
 - Many different pieces of limited software
- Mplus: Integration of methods in one framework
 - Easy to use: Simple language, graphics
 - Powerful: General modeling capabilities

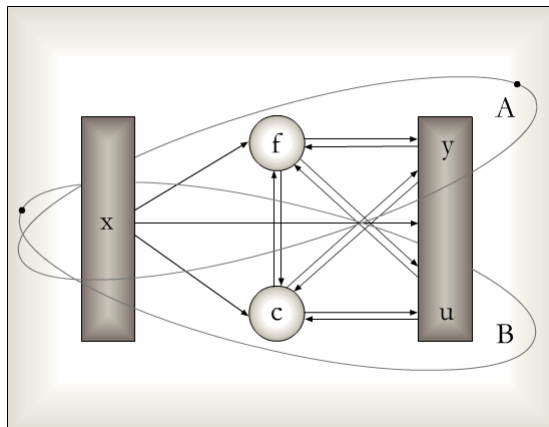
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The General Modeling Framework of Mplus

- Achieves its flexibility from using a combination of categorical and continuous latent variables
- Special cases:
 - Factor analysis and structural equation modeling
 - Growth modeling
 - Mixture (latent class) modeling
 - Multilevel modeling
 - Combinations

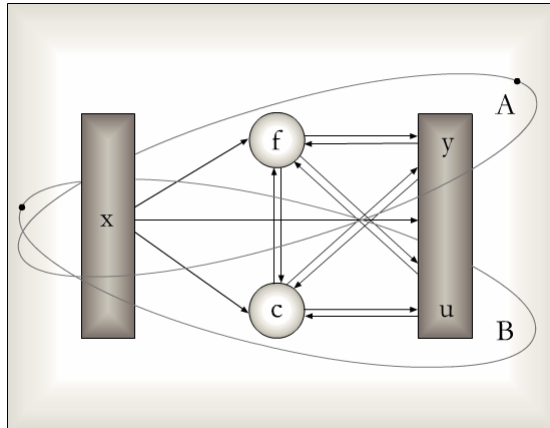
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General Latent Variable Modeling Framework



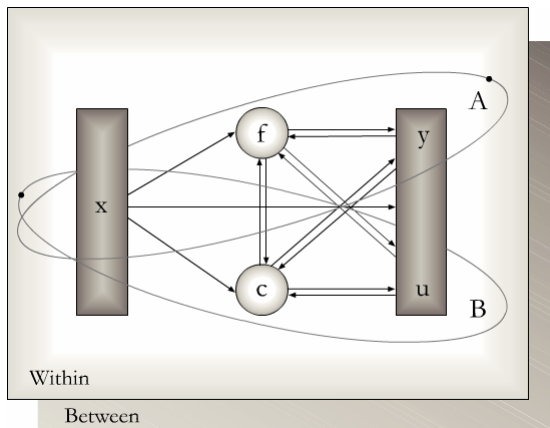
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General Latent Variable Modeling Framework



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General Latent Variable Modeling Framework



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General Latent Variable Modeling Framework

- Muthén, B. (2002). Beyond SEM: General latent variable modeling. *Behaviormetrika*, 29, 81-117
- Asparouhov & Muthen (2004). Maximum-likelihood estimation in general latent variable modeling
- Muthen & Muthen (1998-2004). Mplus Version 3
- Mplus team: Linda Muthen, Bengt Muthen, Tihomir Asparouhov, Thuy Nguyen, Michelle Conn (see www.statmodel.com)

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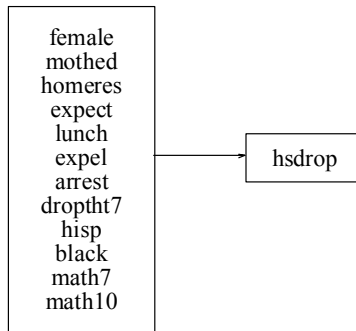
Continuous Latent Variables

- Factor analysis, structural equation modeling
 - Constructs measured with multiple indicators
- Growth modeling
 - Growth factors, random effects: random intercepts and random slopes representing individual differences of development over time (unobserved heterogeneity)
- Survival analysis
 - Frailties
- Missing data modeling

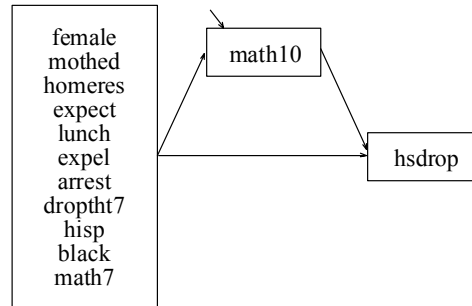
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Path Analysis with a Categorical Outcome and Missing Data on a Mediator

Logistic Regression



Path Analysis



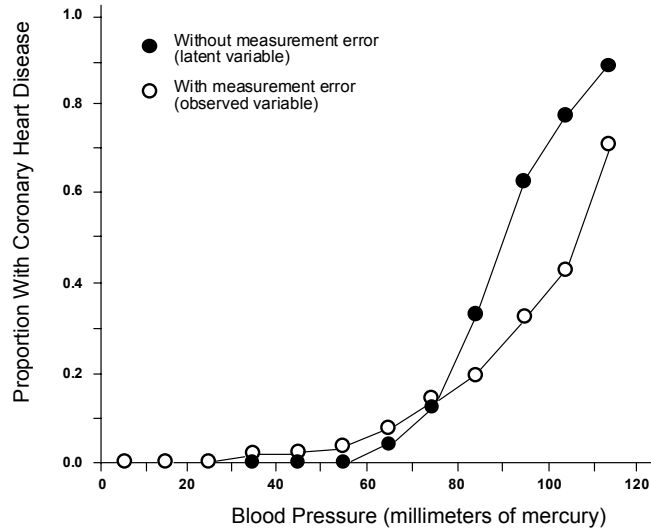
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Continuous Latent Variables: Two Examples

- Muthen (1992). Latent variable modeling in epidemiology. Alcohol Health & Research World, 16, 286-292
 - Blood pressure predicting coronary heart disease
- Nurses' Health Study (Rosner, Willet & Spiegelman, 1989). Nutritional study of 89,538 women.
 - Dietary fat intake questionnaire for everyone
 - Dietary diary for 173 women for 4 1-week periods at 3-month intervals

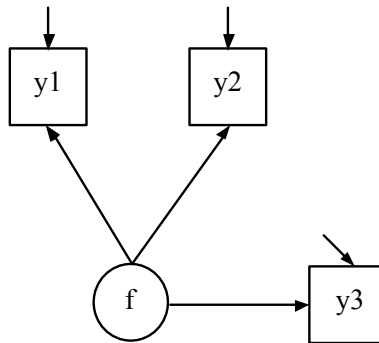
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Measurement Error in a Covariate



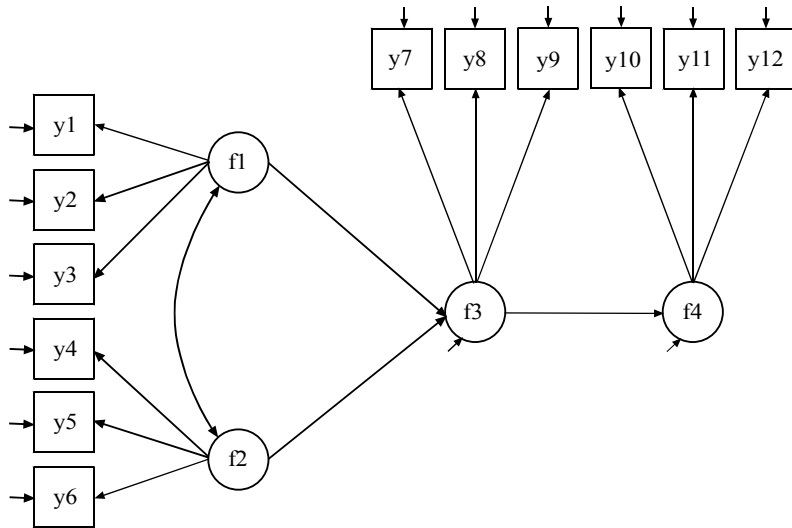
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Measurement Error in a Covariate



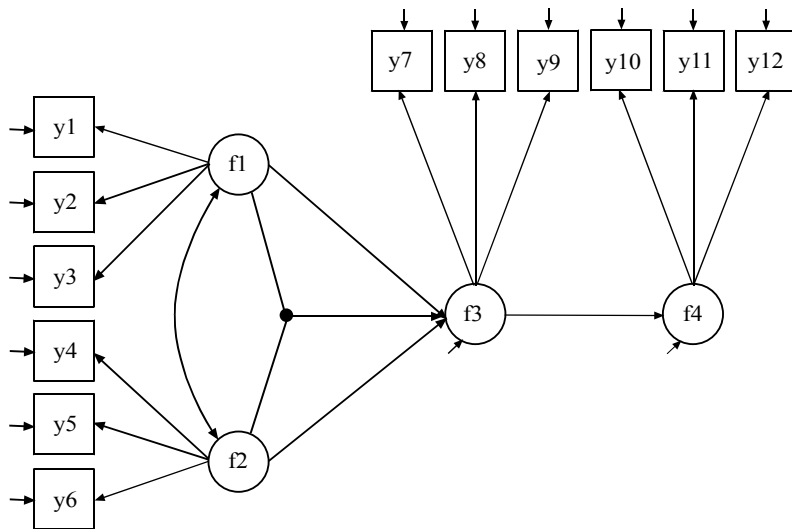
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Structural Equation Model



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Structural Equation Model with Interaction between Latent Variables



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Antisocial Behavior (ASB) Data

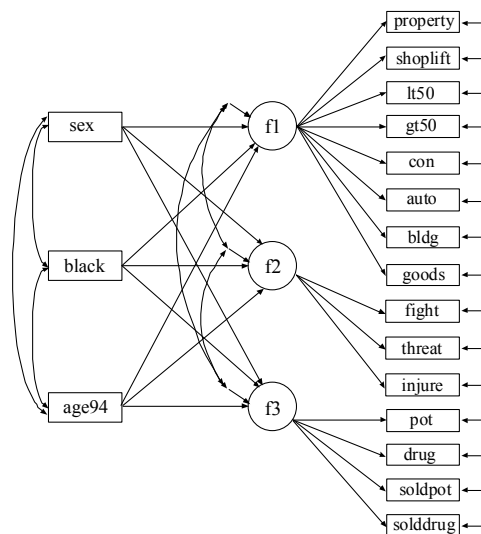
The antisocial behavior (ASB) data were taken from the National Longitudinal Survey of Youth (NLSY) that is sponsored by the Bureau of Labor Statistics. These data are made available to the public by Ohio State University. The data were obtained as a multistage probability sample with oversampling of blacks, Hispanics, and economically disadvantaged non-blacks and non-Hispanics.

Data for the analysis include 15 of the 17 antisocial behavior items that were collected in 1980 when respondents were between the ages of 16 and 23 and the background variables of age, gender, and ethnicity. The ASB items assessed the frequency of various behaviors during the past year. A sample of 7,326 respondents has complete data on the antisocial behavior items and the background variables of age, gender, and ethnicity. Following is a list of the 15 items:

- | | |
|--------------------|----------------------|
| Damaged property | Use other drugs |
| Fighting | Sold marijuana |
| Shoplifting | Sold hard drugs |
| Stole < \$50 | “Con” someone |
| Stole > \$50 | Take auto |
| Seriously threaten | Broken into building |
| Intent to injure | Held stolen goods |
| Use marijuana | |

These items were dichotomized 0/1 with 0 representing never in the last year. An EFA suggested three factors: property offense, person offense, and drug offense.

ASB CFA With Covariates



Input For CFA With Covariates With Categorical Outcomes For 15 ASB Items

TITLE: CFA with covariates with categorical outcomes using
15 antisocial behavior items and 3 covariates

DATA: FILE IS asb.dat;
FORMAT IS 34X 54F2.0;

VARIABLE: NAMES ARE property fight shoplift lt50 gt50 force
threat injure pot drug soldpot solddrug con auto bldg
goods gambling dsm1-dsm22 sex black hisp single
divorce dropout college onset fhist1 fhist2 fhist3
age94 cohort dep abuse;

USEV ARE property-gt50 threat-goods sex black age94

CATEGORICAL ARE property-goods;

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Input For CFA With Covariates With Categorical Outcomes For 15 ASB Items (Continued)

ANALYSIS: TYPE = MEANSTRUCTURE;

MODEL: f1 BY property shoplift-gt50 con-goods;
f2 BY fight threat injure;
f3 BY pot-solddrug;
f1-f3 ON sex black age94;
property-goods ON sex-age94@0;

OUTPUT: STANDARDIZED **TECH2;**

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Output Excerpts CFA With Covariates With Categorical Outcomes For 15 ASB Items

Model Results

		Estimates	S.E.	Est./S.E.	Std	StdYX
F1	BY					
	PROPERTY	1.000	.000	.000	.791	.760
	SHOPLIFT	.974	.023	42.738	.771	.742
	LT50	.915	.023	39.143	.724	.700
	GT50	1.055	.031	33.658	.835	.799
	CON	.752	.024	31.637	.595	.581
	AUTO	.796	.030	26.462	.629	.613
	BLDG	1.084	.030	35.991	.858	.818
	GOODS	1.071	.025	42.697	.847	.809

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Output Excerpts CFA With Covariates With Categorical Outcomes For 15 ASB Items (Continued)

F2	BY					
	FIGHT	1.000	.000	.000	.773	.734
	THREAT	1.096	.035	31.382	.847	.797
	INJURE	1.082	.037	28.888	.836	.787
F3	BY					
	POT	1.000	.000	.000	.866	.851
	DRUG	1.031	.023	45.818	.893	.876
	SOLDPOT	1.046	.023	45.844	.905	.888
	SOLDDRUG	.923	.036	25.684	.799	.787

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Output Excerpts CFA With Covariates With Categorical Outcomes For 15 ASB Items (Continued)

F1	ON					
	SEX	.516	.024	21.206	.653	.326
	BLACK	-.080	.025	-3.168	-.102	-.047
	AGE94	-.054	.006	-9.856	-.069	-.150
F2	ON					
	SEX	.561	.026	21.715	.726	.363
	BLACK	.174	.025	7.087	.225	.103
	AGE94	-.068	.006	-12.286	-.087	-.191
F3	ON					
	SEX	.229	.026	8.760	.265	.132
	BLACK	-.272	.029	-9.384	-.315	-.144
	AGE94	.039	.006	6.481	.045	.099

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Output Excerpts CFA With Covariates With Categorical Outcomes For 15 ASB Items (Continued)

Tests Of Model Fit

Chi-square Test of Model Fit		
Value		1225.266*
Degrees of Freedom		105**
P-Value		0.0000
CFI / TLI		
CFI		0.945
TLI		0.964
RMSEA (Root Mean Square Error Of Approximation)		
Estimate		0.038
WRMR (Weighted Root Mean Square Residual)		
Value		2.498

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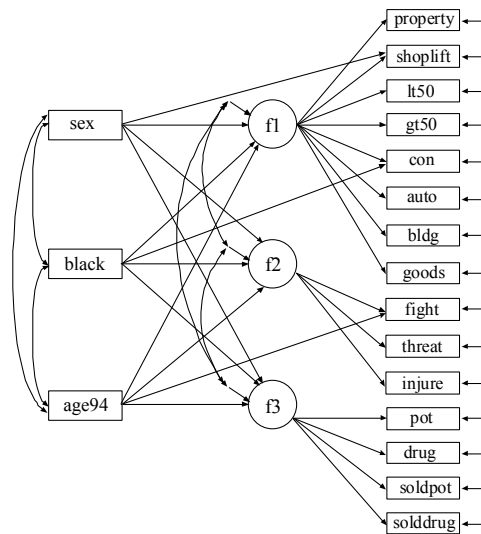
Output Excerpts (Continued)

Tech2 Derivatives With Respect To Gamma

	SEX	BLACK	AGE94
F1	.000	.000	.000
F2	.000	.000	.000
F3	.000	.000	.000
PROPERTY	-.019	.006	.072
FIGHT	-.023	-.015	.109
SHOPLIFT	.039	.001	.003
LT50	-.001	.014	-.072
GT50	-.007	-.008	-.026
THREAT	.009	.015	-.026
ONJURE	.012	-.001	-.074
POT	.011	-.010	-.058
DRUG	.012	.016	-.016
SOLDPOT	-.019	-.003	.060
SOLDDRUG	-.003	-.004	.013
CON	.020	-.030	.051
AUTO	.002	.003	.020
BLDG	-.012	.005	-.003
GOODS	-.013	.003	-.030

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ASB Model With Direct Effects



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Summary of Analysis Results For ASB CFA With Covariates And Direct Effects

Input Specification

```
MODEL:      f1 BY property shoplift-gt50 con-goods;
            f2 BY fight threat injure;
            f3 BY pot-solddrug;

            f1-f3 ON sex black age94;

            shoplift ON sex;
            con ON black;
            fight ON age94;
```

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Summary of Analysis Results For ASB CFA With Covariates And Direct Effects

Test Of Model Fit

Chi-square Test of Model Fit	
Value	946.256*
Degrees of Freedom	102**
P-Value	0.0000
CFI/TLI	
CFI	0.959
TLI	0.972
RMSEA (Root Mean Square Error Of Approximation)	
Estimate	0.034
WRMR (Weighted Root Mean Square Residual)	
Value	2.198

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Summary of Analysis Results For ASB CFA With Covariates And Direct Effects (Continued)

		Estimates	S.E.	Est./S.E.	Std	StdYX
F1	BY					
	SHOPLIFT	1.002	.024	42.183	.805	.793
F1	ON					
	SEX	.596	.026	22.958	.742	.371
SHOPLIFT	ON					
	SEX	-.385	.033	-11.594	-.385	-.190
CON	ON					
	BLACK	.305	.034	8.929	.305	.136
FIGHT	ON					
	AGE94	-.068	.008	-8.467	-.068	-.138
Thresholds						
	SHOPLIFT\$1	.558	.033	17.015	.558	.558
R-SQUARE						
	Observed	Residual				
	Variable	Variance	R-Square			
	SHOPLIFT	.461	.552			27

Interpretation of Direct Effects

- Look at indirect effect of covariate on factor indicator via the factor
- Look at the direct effect of covariate on factor indicator

Shoplift On Gender

- Indirect effect of gender on shoplift
 - F1 has a positive relationship with gender – males have a higher mean than females on the f1 factor
 - Shoplift has a positive loading on the f1 factor
 - Conclusion: males are expected to have a higher probability of shoplifting
- Effect of gender on shoplift
 - Direct effect is negative – for a given factor value, males have a lower probability of shoplifting than females
 - Conclusion – shoplift is not invariant

Calculating Item Probabilities

The model with a direct effect from x to item y_j ,

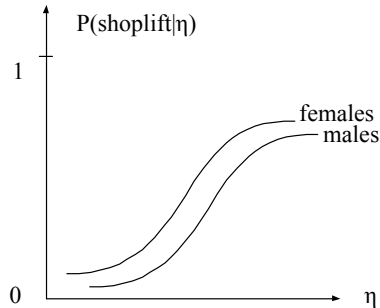
$$y_{ij}^* = \lambda_j \eta_i + \kappa_j x_i + \varepsilon_{ij}, \quad (45)$$

gives the conditional probability of a $y = 1$ response

$$P(y_{ij} = 1 | \eta_{ij}, x_i) = 1 - F[(\tau_j - \lambda_j \eta_i - \kappa_j x_i) \theta_{jj}^{-1/2}], \quad (46)$$

where F is the normal distribution function.

For example, for the item shoplift, $\tau_j = 0.558$, $\kappa_j = -0.385$, $\theta_{jj} = 0.461$. At $\eta = 0$, the probability is 0.21 for females ($x = 0$) and 0.08 for males ($x = 1$).



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Calculating Item Probabilities Cont'd

Consider

$$P(y_{ij} = 1 | \eta_{ij}, x_i) = 1 - F[(\tau_j - \lambda_j \eta_i - \kappa_j x_i) \theta_{jj}^{-1/2}], \quad (47)$$

Using $\tau_j = 0.558$, $\kappa_j = -0.385$, $\theta_{jj} = 0.461$, and $\eta = 0$.

Here, $\theta_{jj}^{-1/2} = \frac{1}{\sqrt{\theta_{jj}}} = \frac{1}{\sqrt{0.461}} = 1.473$.

For females ($x = 0$):

1. $(\tau_j - \lambda_j \eta_i - \kappa_j x_i) = 0.558 - 1.002 \times 0 - (-0.385) \times 0 = 0.558$.
2. $(\tau_j - \lambda_j \eta_i - \kappa_j x_i) \theta_{jj}^{-1/2} = 0.558 \times 1.473 = 0.822$.
3. $F[0.822] = 0.794$ using a z table
4. $1 - 0.794 = 0.206$.

For males ($x = 1$):

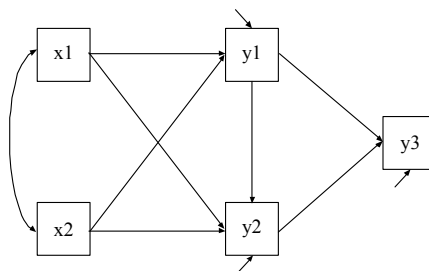
1. $(\tau_j - \lambda_j \eta_i - \kappa_j x_i) = 0.558 - 1.002 \times 0 - (-0.385) \times 1 = 0.943$.
2. $(\tau_j - \lambda_j \eta_i - \kappa_j x_i) \theta_{jj}^{-1/2} = 0.943 \times 1.473 = 1.389$.
3. $F[1.389] = 0.918$ using a z table.
4. $1 - 0.918 = 0.082$.

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New SEM Features In Mplus Version 3

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The MODEL INDIRECT Command



MODEL INDIRECT has two options:

- IND – used to request a specific indirect effect or a set of indirect effects
- VIA – used to request a set of indirect effects that includes specific mediators

MODEL INDIRECT

```
y3 IND y1 x1;      ! x1 -> y1 -> y3
y3 IND y2 x2;      ! x2 -> y2 -> y3
y3 IND x1;          ! x1 -> y1 -> y3
                   ! x1 -> y2 -> y3
                   ! x1 -> y1 -> y2 -> y3
y3 VIA y2 x1;      ! x1 -> y2 -> y3
                   ! x1 -> y1 -> y2 -> y3
```

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The MODEL INDIRECT Command

MODEL INDIRECT is used to request indirect effects and their standard errors. Delta method standard errors are computed as the default.

The STANDARDIZED option of the OUTPUT command can be used to obtain standardized indirect effects.

The BOOTSTRAP option of the ANALYSIS command can be used to obtain bootstrap standard errors for the indirect effects.

The CINTERVAL option of the OUTPUT command can be used to obtain confidence intervals for the indirect effects and the standardized indirect effects. Three types of 95% and 99% confidence intervals can be obtained: symmetric, bootstrap, or bias-corrected bootstrap confidence intervals. The bootstrapped distribution of each parameter estimate is used to determine the bootstrap and bias-corrected bootstrap confidence intervals. These intervals take non-normality of the parameter estimate distribution into account. As a result, they are not necessarily symmetric around the parameter estimate.

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The MODEL CONSTRAINT Command

MODEL CONSTRAINT is used to define linear and non-linear constraints on the parameters in the model. All functions available in the DEFINE command are available for linear and non-linear constraints. Parameters in the model are given labels by placing a name in parentheses after the parameter.

```
MODEL: y ON x1 (p1)
        x2 (p2)
        x3 (p3);
```

```
MODEL CONSTRAINT:
    p1 = p2**2 + p3**2;
```

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Interaction Modeling Using ML For Observed And Latent Variables

Types of Variables	Interaction Options
observed continuous with observed continuous	DEFINE
observed categorical with observed continuous	DEFINE Multiple Group
observed continuous with continuous latent	XWITH
observed categorical with continuous latent	XWITH Multiple Group
observed continuous with categorical latent	MIXTURE
observed categorical with categorical latent	MIXTURE KNOWNCLASS
continuous latent with continuous latent	XWITH
continuous latent with categorical latent	MIXTURE
categorical latent with categorical latent	MIXTURE

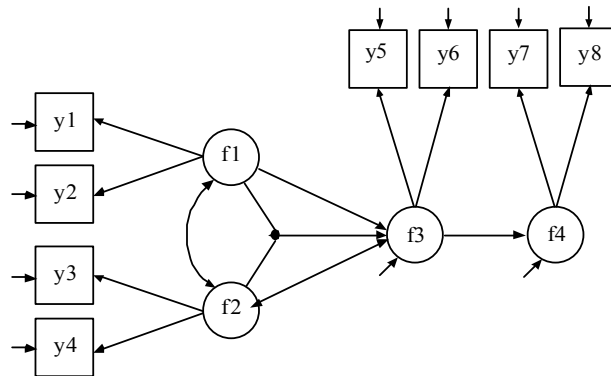
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The XWITH Option Of The MODEL Command

The XWITH option is used with TYPE=RANDOM to define interactions between continuous latent variables or between continuous latent variables and observed variables. XWITH is short for multiplied with. It is used in conjunction with the | symbol to name and define interaction variables in a model. Following is an example of how to use XWITH and the | symbol to name and define an interaction:

```
f1f2 | f1 XWITH f2;  
f1y | f1 XWITH y;
```

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Input For An SEM Model With An Interaction Between Two Latent Variables

```

TITLE:      this an example of a structural equation model with
            an
            interaction between two latent variables
DATA:       FILE = firstSEMInter.dat;
VARIABLE:   NAMES = y1-y8;
ANALYSIS:   TYPE = RANDOM;
            ALGORITHM = INTEGRATION;
MODEL:      f1 BY y1 y2;
            f2 BY y3 y4;
            f3 BY y5 y6;
            f4 BY y7 y8;

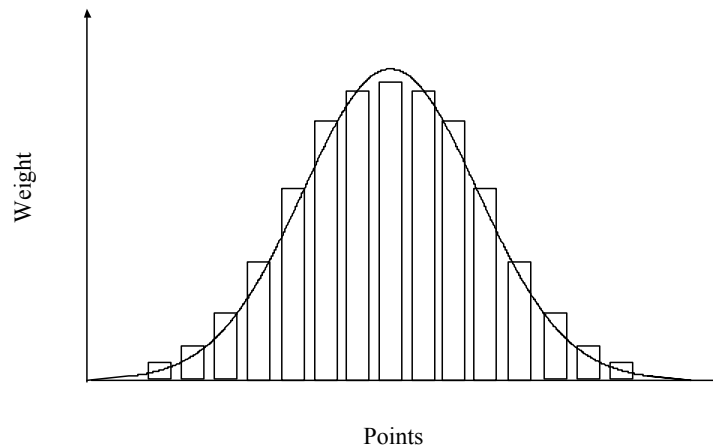
            f4 ON f3;
            f3 ON f1 f2;

            f1f2 | f1 XWITH f2;

            f3 ON f1f2;
OUTPUT:     TECH8;
  
```

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Numerical Integration



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Numerical Integration

Numerical integration is needed with maximum likelihood estimation when the posterior distribution for the latent variables does not have a closed form expression. This occurs for models with categorical outcomes that are influenced by continuous latent variables, for models with interactions involving continuous latent variables, and for certain models with random slopes such as multilevel mixture models.

When the posterior distribution does not have a closed form, it is necessary to integrate over the density of the latent variables multiplied by the conditional distribution of the outcomes given the latent variables. Numerical integration approximates this integration by using a weighted sum over a set of integration points (quadrature nodes) representing values of the latent variable.

Numerical integration is computationally heavy and thereby time-consuming because the integration must be done at each iteration, both when computing the function value and when computing the derivative values. The computational burden increases as a function of the number of integration points, increases linearly as a function of the number of observations, and increases exponentially as a function of the dimension of integration, that is, the number of latent variables for which numerical integration is needed.

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Practical Aspects of Numerical Integration

- Types of numerical integration available in Mplus with or without adaptive quadrature
 - Standard (rectangular, trapezoid) – default with 15 integration points per dimension
 - Gauss-Hermite
 - Monte Carlo
 - Computational burden for latent variables that need numerical integration
 - One or two latent variables Light
 - Three to five latent variables Heavy
 - Over five latent variables Very Heavy
- Suggestions for using numerical integration
- Start with a model with a small number of random effects and add more one at a time
 - Start with an analysis with TECH8 and ITERATIONS=1 to obtain information from the screen printing on the dimensions of integration and the time required for one iteration and with TECH1 to check model specifications
 - With more than 3 dimensions, reduce the number of integration points to 10 or use Monte Carlo integration with the default of 500 integration points
 - If the TECH8 output shows large negative values in the column labeled ABS CHANGE, increase the number of integration points to improve the precision of the numerical integration and resolve convergence problems.

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Numerical Integration Theory

Maximum likelihood estimation using the EM algorithm computes in each iteration the posterior distribution for normally distributed latent variables f ,

$$[f|y] = [f][y|f] / [y], \quad (97)$$

where the marginal density for $[y]$ is expressed by integration

$$[y] = \int [f][y|f] df. \quad (98)$$

- Numerical integration is not needed: Normally distributed y – the posterior distribution is normal
- Numerical integration is needed:
 - Categorical outcomes u influenced by continuous latent variables f , because $[u]$ has no closed form
 - Latent variable interactions $f \times x, f \times y, f_1 \times f_2$, where $[y]$ has no closed form, for example

$$[y] = \int [f_1, f_2][y|f_1, f_2, f_1 f_2] df_1 df_2 \quad (99)$$
 - Random slopes, e.g. with two-level mixture modeling

Numerical integration approximates the integral by a sum

$$[y] = \int [f][y|f] df = \sum_{k=1}^K w_k [y|f_k] \quad (100)$$

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General Latent Variable Modeling Using Mplus Version 3

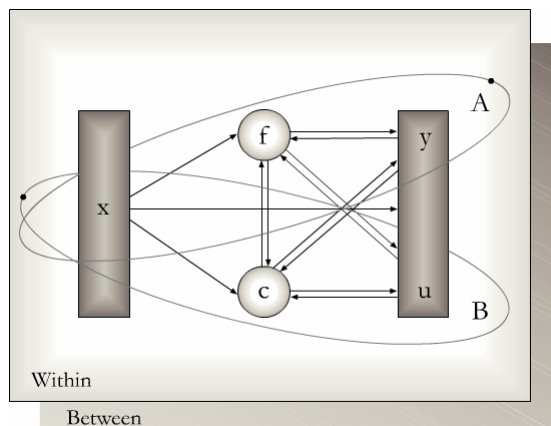
Block 2: Growth Modeling

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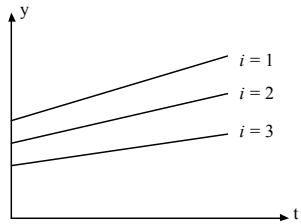
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General Latent Variable Modeling Framework



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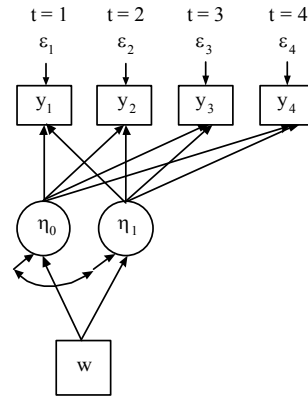
Individual Development Over Time



$$(1) \quad y_{it} = \eta_{0i} + \eta_{1i} x_t + \varepsilon_{it}$$

$$(2a) \quad \eta_{0i} = \alpha_0 + \gamma_0 w_i + \zeta_{0i}$$

$$(2b) \quad \eta_{1i} = \alpha_1 + \gamma_1 w_i + \zeta_{1i}$$

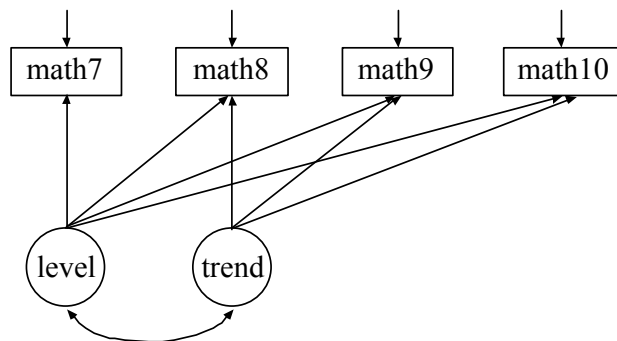


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Advantages of Growth Modeling in a Latent Variable Framework

- Flexible curve shape
- Individually-varying times of observation
- Random effects (intercepts, slopes) integrated with other latent variables
- Regressions among random effects
- Multiple processes
- Multiple populations
- Multiple indicators
- Embedded growth models
- Categorical latent variables: growth mixtures

4



5

Input For LSAY Linear Growth Model Without Covariates

```

TITLE:    LSAY For Younger Females With Listwise Deletion
          Linear Growth Model Without Covariates

DATA:     FILE IS lsay.dat;
          FORMAT IS 3F8.0 F8.4 8F8.2 3F8.0;

VARIABLE: NAMES ARE cohort id school weight math7 math8 math9
          math10 att7 att8 att9 att10 gender mothed homeres;
          USEOBS = (gender EQ 1 AND cohort EQ 2);
          MISSING = ALL (999);
          USEVAR = math7-math10;
  
```

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Input For LSAY Linear Growth Model Without Covariates (Continued)

```
ANALYSIS:      TYPE = MEANSTRUCTURE;  
  
MODEL:    level BY math7-math10@1;  
          trend BY math7@0 math8@1 math9@2 math10@3;  
          [math7-math10@0];  
          [level trend];  
  
OUTPUT:    SAMPSTAT STANDARDIZED MODINDICES (3.84);  
  
!New Version 3 Language For Growth Models  
!MODEL: level trend | math7@0 math8@1 math9@2 math10@3
```

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Output Excerpts LSAY Linear Growth Model Without Covariates

Tests of Model Fit

Chi-Square Test of Model Fit

Value	22.664
Degrees of Freedom	5
P-Value	0.0004

CFI/TLI

CFI	0.995
TLI	0.994

RMSEA (Root Mean Square Error Of Approximation)

Estimate	0.060
90 Percent C.I.	0.036 0.086
Probability RMSEA <= .05	0.223

SRMR (Standardized Root Mean Square Residual)

Value	0.025
-------	-------

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Output Excerpts LSAY Linear Growth Model Without Covariates (Continued)

Modification Indices

		M.I.	E.P.C.	Std.E.P.C.	StdYX E.P.C.
TREND	BY MATH7	6.793	0.185	0.254	0.029
TREND	BY MATH8	14.694	-0.169	-0.233	-0.025
TREND	BY MATH9	9.766	0.155	0.213	0.021

Output Excerpts LSAY Linear Growth Model Without Covariates

Model Results

		Estimates	S.E.	Est./S.E.	Std	StdYX
LEVEL	BY					
	MATH7	1.000	.000	.000	8.029	.906
	MATH8	1.000	.000	.000	8.029	.861
	MATH9	1.000	.000	.000	8.029	.800
	MATH10	1.000	.000	.000	8.029	.708
TREND	BY					
	MATH7	.000	.000	.000	.000	.000
	MATH8	1.000	.000	.000	1.377	.148
	MATH9	2.000	.000	.000	2.753	.274
	MATH10	3.000	.000	.000	4.130	.364

Output Excerpts LSAY Linear Growth Model Without Covariates (Continued)

LEVEL	WITH					
	TREND	3.491	.730	4.780	.316	.316
Residual Variances						
	MATH7	14.105	1.253	11.259	14.105	.180
	MATH8	13.525	.866	15.610	13.525	.156
	MATH9	14.726	.989	14.897	14.726	.146
	MATH10	25.989	1.870	13.898	25.989	.202
Variances						
	LEVEL	64.469	3.428	18.809	1.000	1.000
	TREND	1.895	.322	5.894	1.000	1.000

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Output Excerpts LSAY Linear Growth Model Without Covariates (Continued)

Means						
	LEVEL	52.623	.275	191.076	6.554	6.554
	TREND	3.105	.075	41.210	2.255	2.255
Intercepts						
	MATH7	.000	.000	.000	.000	.000
	MATH8	.000	.000	.000	.000	.000
	MATH9	.000	.000	.000	.000	.000
	MATH10	.000	.000	.000	.000	.000
R-Square						
	Observed					
	Variable	R-Square				
	MATH7	0.820				
	MATH8	0.844				
	MATH9	0.854				
	MATH10	0.798				

12

Growth Model With Individually Varying Times Of Observation And Random Slopes For Time-Varying Covariates

13

Growth Modeling In Multilevel Terms

Time point t , individual i (two-level modeling, no clustering):

- y_{ti} : repeated measures on the outcome, e.g. math achievement
- a_{1ti} : time-related variable (time scores); e.g. grade 7-10
- a_{2ti} : time-varying covariate, e.g. math course taking
- x_i : time-invariant covariate, e.g. grade 7 expectations

Two-level analysis with individually-varying times of observation and random slopes for time-varying covariates:

$$\text{Level 1: } y_{ti} = \pi_{0i} + \pi_{1i} a_{1ti} + \pi_{2i} a_{2ti} + e_{ti}, \quad (55)$$

$$\text{Level 2: } \begin{cases} \pi_{0i} = \beta_{00} + \beta_{01} x_i + r_{0i}, \\ \pi_{1i} = \beta_{10} + \beta_{11} x_i + r_{1i}, \\ \pi_{2i} = \beta_{20} + \beta_{21} x_i + r_{2i}. \end{cases} \quad (56)$$

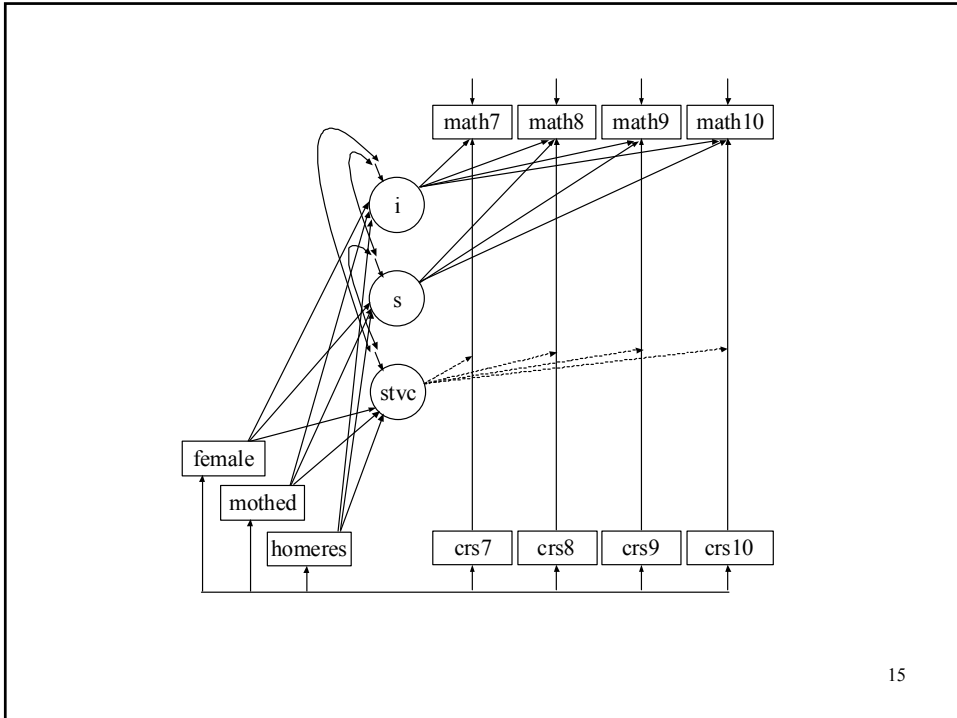
Time scores a_{1ti} read in as data (not loading parameters).

- π_{2i} possible with time-varying random slope variances
- Flexible correlation structure for $V(e) = \Theta(T \times T)$
- Regressions among random coefficients possible, e.g.

$$\pi_{1i} = \beta_{10} + \gamma_1 \pi_{0i} + \beta_{11} x_i + r_{1i}, \quad (57)$$

$$\pi_{2i} = \beta_{20} + \gamma_2 \pi_{0i} + \beta_{21} x_i + r_{2i}. \quad (58)$$

14



Input For Growth Model With Individually Varying Times Of Observation

```

TITLE:    growth model with individually varying times of
          observation and random slopes

DATA:     FILE IS lsaynew.dat;

VARIABLE: NAMES ARE math7 math8 math9 math10 crs7 crs8 crs9
          crs10 female mothed homeres a7-a10;

          ! crs7-crs10 = highest math course taken during each
          ! grade (0=no course, 1=low,basic, 2=average, 3=high,
          ! 4=pre-algebra, 5=algebra I, 6=geometry,
          ! 7=algebra II, 8=pre-calc, 9=calculus)

MISSING ARE ALL(9999);
CENTER = GRANDMEAN(crs7-crs10 mothed homeres);
TSCORES = a7-a10;

DEFINE:   math7 = math7/10;
          math8 = math8/10;
          math9 = math9/10;
          math10 = math10/10;

```

16

Input For Growth Model With Individually Varying Times Of Observation (Continued)

```
ANALYSIS: TYPE = RANDOM MISSING;
           ESTIMATOR = ML;
           MCONVERGENCE = .001;

MODEL:    i s | math7-math10 AT a7-a10;

           stvc | math7 ON crs7;
           stvc | math8 ON crs8;
           stvc | math9 ON crs9;
           stvc | math10 ON crs10;

           i ON female mothed homeres;
           s ON female mothed homeres;
           stvc on female mothed homeres;

           i WITH s;
           stvc WITH i;
           stvc WITH s;

OUTPUT:   TECH8;
```

17

Output Excerpts For Growth Model With Individually Varying Times Of Observation And Random Slopes For Time-Varying Covariates

n = 2271

Tests of Model Fit

Loglikelihood

H0 Value	-8199.311
----------	-----------

Information Criteria

Number of Free Parameters	22
Akaike (AIC)	16442.623
Bayesian (BIC)	16568.638
Sample-Size Adjusted BIC	16498.740
(n* = (n + 2) / 24)	

18

**Output Excerpts For Growth Model With Individually Varying Times Of
Observation And Random Slopes For Time-Varying Covariates (Continued)**

		Estimates	S.E.	Est./S.E.
I	ON			
	FEMALE	0.187	0.036	5.247
	MOTHED	0.187	0.018	10.231
	HOMERES	0.159	0.011	14.194
S	ON			
	FEMALE	-0.025	0.012	-2.017
	MOTHED	0.015	0.006	2.429
	HOMERES	0.019	0.004	4.835
STVC	ON			
	FEMALE	-0.008	0.013	-0.590
	MOTHED	0.003	0.007	0.429
	HOMERES	0.009	0.004	2.167
I	WITH			
S		0.038	0.006	6.445
STVC	WITH			
I		0.011	0.005	2.087
S		0.004	0.002	2.033

19

**Output Excerpts For Growth Model With Individually Varying Times Of
Observation And Random Slopes For Time-Varying Covariates (Continued)**

Intercepts				
	MTH7	0.000	0.000	0.000
	MTH8	0.000	0.000	0.000
	MTH9	0.000	0.000	0.000
	MTH10	0.000	0.000	0.000
	I	4.992	0.025	198.456
	S	0.417	0.009	47.275
	STVC	0.113	0.010	11.416
Residual Variances				
	MTH7	0.185	0.011	16.464
	MTH8	0.178	0.008	22.232
	MTH9	0.156	0.008	18.497
	MTH10	0.169	0.014	12.500
	I	0.570	0.023	25.087
	S	0.036	0.003	12.064
	STVC	0.012	0.002	5.055

20

Random Slopes

- In single-level modeling random slopes β_i describe variation across individuals i ,

$$y_i = \alpha_i + \beta_i x_i + \varepsilon_i, \quad (100)$$

$$\alpha_i = \alpha + \zeta_{0i}, \quad (101)$$

$$\beta_i = \beta + \zeta_{1i}, \quad (102)$$

Resulting in heteroscedastic residual variances

$$V(y_i | x_i) = V(\beta_i) x_i^2 + \theta \quad (103)$$

- In two-level modeling random slopes β_j describe variation across clusters j

$$y_{ij} = a_j + \beta_j x_{ij} + \varepsilon_{ij}, \quad (104)$$

$$a_j = a + \zeta_{0j}, \quad (105)$$

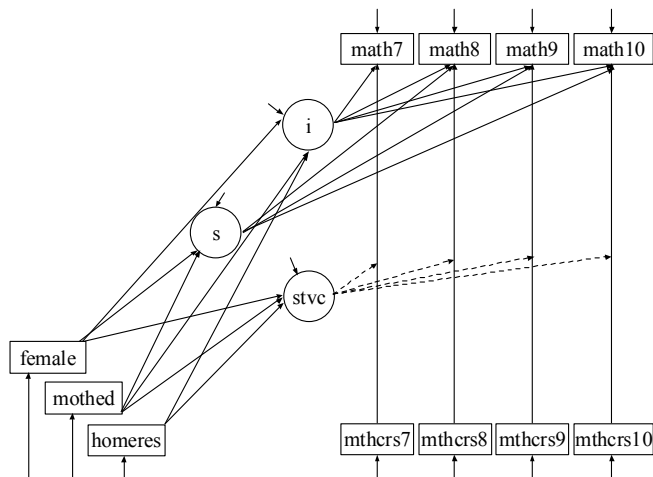
$$\beta_j = \beta + \zeta_{1j}, \quad (106)$$

A small variance for a random slope typically leads to slow convergence of the ML-EM iterations. This suggests respecifying the slope as fixed.

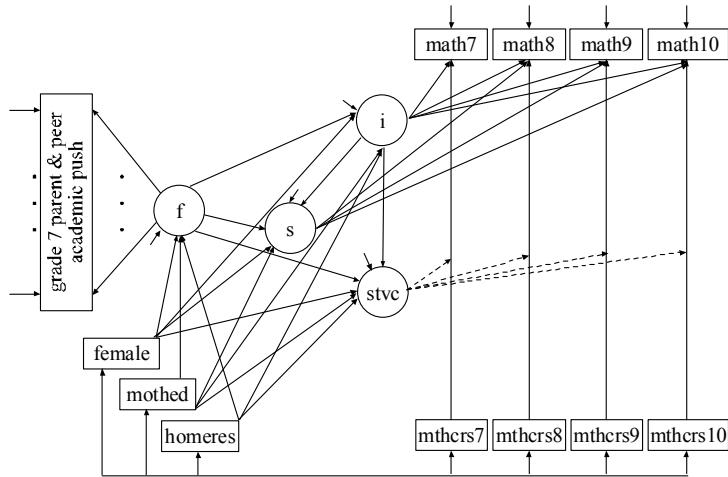
Mplus allows random slopes for predictors that are

- Observed covariates
- Observed dependent variables (Version 3)
- Continuous latent variables (Version 3)

Growth Modeling with Time-Varying Covariates

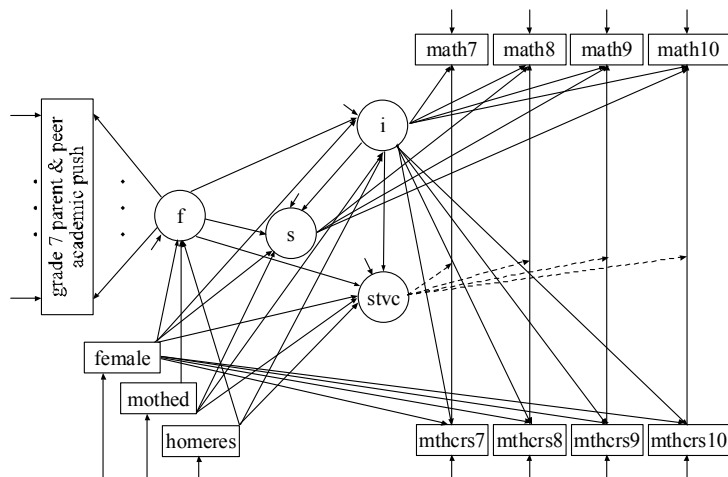


A Generalized Growth Model



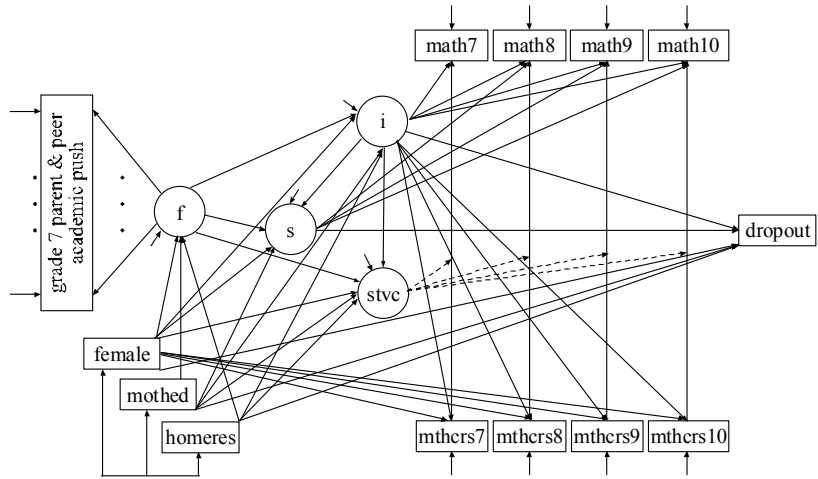
23

A Generalized Growth Model



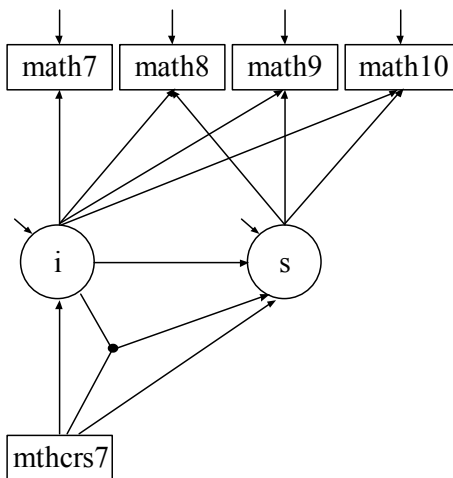
24

A Generalized Growth Model



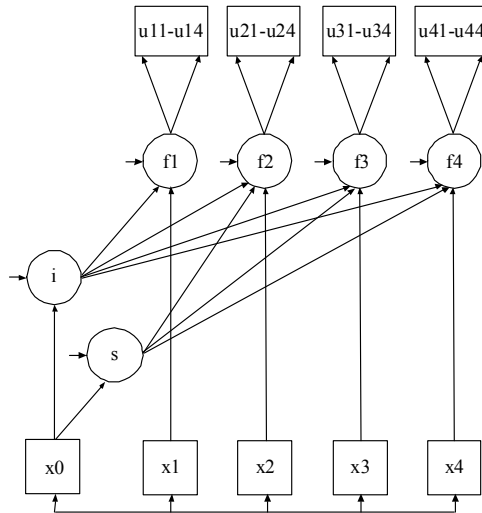
25

Growth Modeling with a Latent Variable Interaction



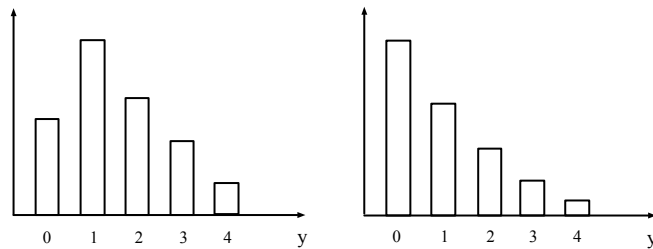
26

Growth in Factors Measured by Multiple Categorical Indicators



27

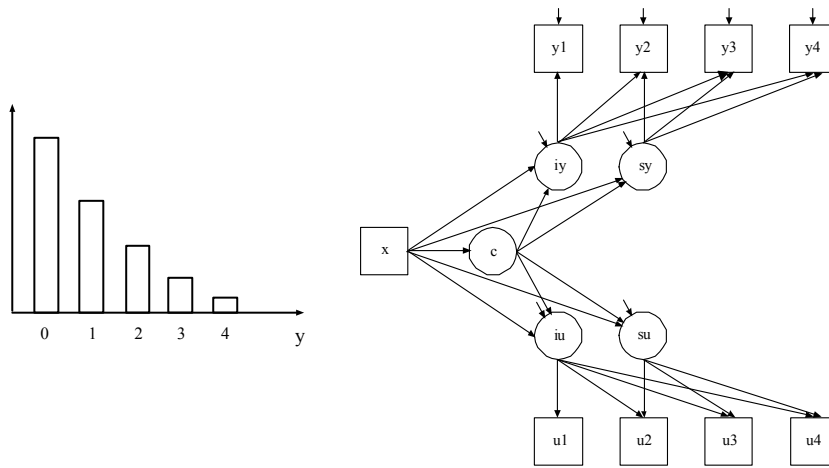
Modeling with a Preponderance of Zeros



- Outcomes: non-normal continuous – count – categorical
- Censored-normal modeling
- Two-part (semicontinuous modeling): Duan et al. (1983), Olsen & Shafer (2001)
- Mixture models, e.g. zero-inflated (mixture) Poisson (Roeder et al., 1999), censored-inflated, mover-stayer latent transition models, growth mixture models
- Onset (survival) followed by growth: Albert & Shih (2003)

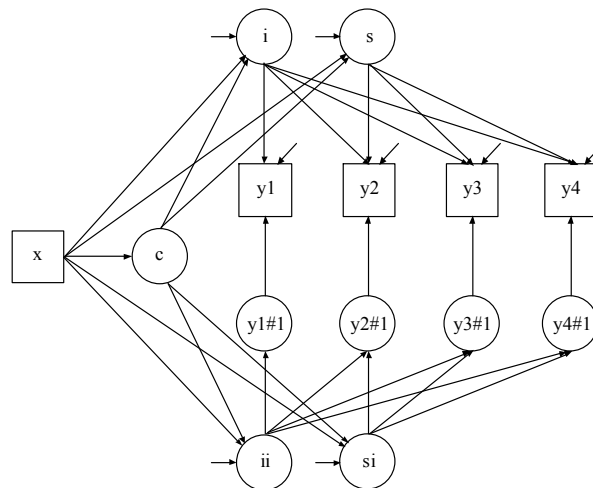
28

Two-Part (Semicontinuous) Growth Modeling



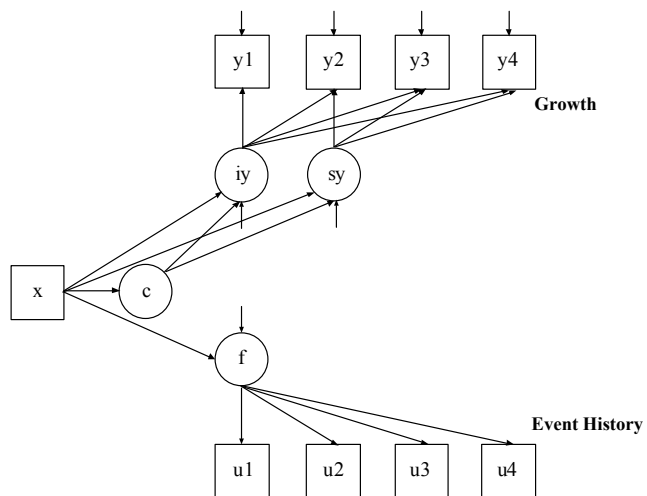
29

Inflated Growth Modeling (Two Classes At Each Time Point)



30

Onset (Survival) Followed by Growth



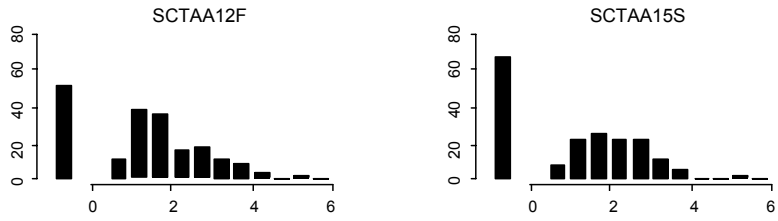
31

Two-Part Growth Modeling

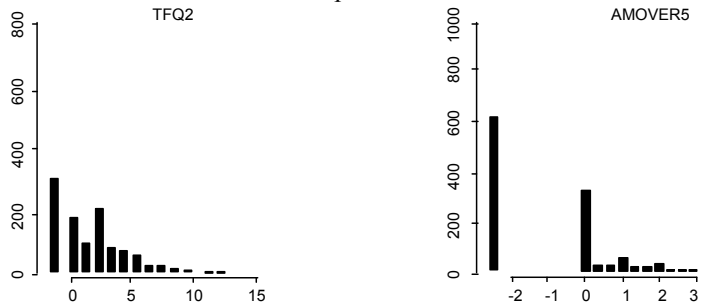
32

Two Types of Skewed Distributions

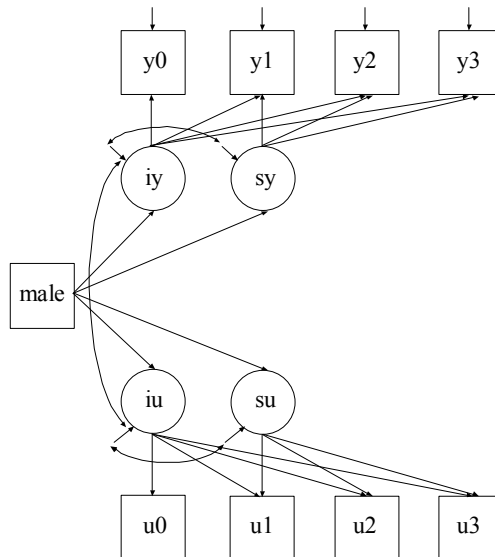
Normal mixture components



Preponderance of zeroes



33



34

Input For Step 1 Of A Two-Part Growth Model

```
TITLE:      step 1 of a two-part growth model
            Amover u    y
            >0    1    >0
            0    0    999
            999 999 999

DATA:      FILE = amp.dat;

VARIABLE:  NAMES ARE caseid
            amover0 ovrdrnk0 illdrnk0 vrydrn0
            amover1 ovrdrnk1 illdrnk1 vrydrn1
            amover2 ovrdrnk2 illdrnk2 vrydrn2
            amover3 ovrdrnk3 illdrnk3 vrydrn3
            amover4 ovrdrnk4 illdrnk4 vrydrn4
            amover5 ovrdrnk5 illdrnk5 vrydrn5
            amover6 ovrdrnk6 illdrnk6 vrydrn6
            tfq0-tfq6 v2 sex race livewith
            agedrnk0-agedrnk6 grades0-grades6;
            USEV = amover0 amover1 amover2 amover3
            sex race u0-u3 y0-y3;
            !MISSING = ALL (999);
```

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Input For Step 1 Of A Two-Part Growth Model (Continued)

```
DEFINE:    u0 = 1;                                !binary part of variable
            IF(amover0 eq 0) THEN u0 = 0;
            IF(amover0 eq 999) THEN u0 = 999;
            y0 = amover0;                          !continuous part of variable
            IF (amover0 eq 0) THEN y0 = 999;
            u1 = 1;
            IF(amover1 eq 0) THEN u1 = 0;
            IF(amover1 eq 999) THEN u1 = 999;
            y1 = amover1;
            IF(amover1 eq 0) THEN y1 = 999;
            u2 = 1;
            IF(amover2 eq 0) THEN u2 = 0;
            IF(amover2 eq 999) THEN u2 = 999;
            y2 = amover2;
            IF(amover2 eq 0) THEN y2 = 999;
            u3 = 1;
            IF(amover3 eq 0) THEN u3 = 0;
            IF(amover3 eq 999) THEN u3 = 999;
            y3 = amover3;
            IF(amover3 eq 0) THEN y3 = 999;

ANALYSIS:  TYPE = BASIC;

SAVEDATA:  FILE = ampyu.dat;
```

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Output Excerpts Step 1 Of A Two-Part Growth Model

SAVEDATA Information

Order and format of variables

```
AMOVER0 F10.3
AMOVER1 F10.3
AMOVER2 F10.3
AMOVER3 F10.3
SEX      F10.3
RACE     F10.3
U0       F10.3
U1       F10.3
U2       F10.3
U3       F10.3
Y0       F10.3
Y1       F10.3
Y2       F10.3
Y3       F10.3
```

Save file

ampyu.dat

Save file format

14F10.3

Save file record length 1000

37

Input For Step 2 Of A Two-Part Growth Model

```
TITLE:      two-part growth model with linear growth for both
             parts
DATA:       FILE = ampya.dat;
VARIABLE:   NAMES = amover0-amover3 sex race u0-u3 y0-y3;
             USEV = u0-u3 y0-y3 male;
             USEOBS = u0 NE 999;
             MISSING = ALL (999);
             CATEGORICAL = u0-u3;
DEFINE:    Male = 2-sex;
```

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Input For Step 2 Of A Two-Part Growth Model (Continued)

```
ANALYSIS:  TYPE = MISSING;
           ESTIMATOR = ML;
           ALGORITHM = INTEGRATION;
           COVERAGE = .09;
MODEL:     iu su | u0@0 u1@0.5 u2@1.5 u3@2.5;
           iy sy | y0@0 y1@0.5 y2@1.5 y3@2.5;
           iu-sy ON male;
           ! estimate the residual covariances
           ! iu with su, iy with sy, and iu with iy
           iu WITH sy@0;
           su WITH iy-sy@0;
OUTPUT:    PATTERNS SAMPSTAT STANDARDIZED TECH1 TECH4 TECH8;
PLOT:      TYPE = PLOT3;
           GROWTH = u0-u3(su) | y0-y3(sy);
```

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Output Excerpts Step 2 Of A Two-Part Growth Model

Tests of Model Fit

Loglikelihood

H0 Value	-3277.101
----------	-----------

Information Criteria

Number of Free parameters	19
Akaike (AIC)	6592.202
Bayesian (BIC)	6689.444
Sample-Size Adjusted BIC	6629.092
(n* = (n + 2) / 24)	

40

Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

Model Results

		Estimates	S.E.	Est./S.E.	Std	StdYX
IU						
	U0	1.000	0.000	0.000	2.839	0.843
	U1	1.000	0.000	0.000	2.839	0.882
	U2	1.000	0.000	0.000	2.839	0.926
	U3	1.000	0.000	0.000	2.839	0.905
SU						
	U0	0.000	0.000	0.000	0.000	0.000
	U1	0.500	0.000	0.000	0.416	0.129
	U2	1.500	0.000	0.000	1.249	0.407
	U3	2.500	0.000	0.000	2.082	0.664

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

IY						
	Y0	1.000	0.000	0.000	0.534	0.787
	Y1	1.000	0.000	0.000	0.534	0.738
	Y2	1.000	0.000	0.000	0.534	0.740
	Y3	1.000	0.000	0.000	0.534	0.644
SY						
	Y0	0.000	0.000	0.000	0.000	0.000
	Y1	0.500	0.000	0.000	0.117	0.162
	Y2	1.500	0.000	0.000	0.351	0.487
	Y3	2.500	0.000	0.000	0.586	0.707

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

IU	ON					
	MALE	0.569	0.234	2.433	0.200	0.100
SU	ON					
	MALE	-0.181	0.119	-1.518	-0.218	-0.109
IY	ON					
	MALE	0.149	0.061	2.456	0.279	0.139
SY	ON					
	MALE	-0.068	0.038	-1.790	-0.290	-0.145
IU	WITH					
	SU	-1.144	0.326	-3.509	-0.484	-0.484
	IY	1.193	0.134	8.897	0.788	0.788
	SY	0.000	0.000	0.000	0.000	0.000
IY	WITH					
	SY	-0.039	0.019	-2.109	-0.316	-0.316
SU	WITH					
	IY	0.000	0.000	0.000	0.000	0.000
	SY	0.000	0.000	0.000	0.000	0.000

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

Intercepts						
Y0		0.000	0.000	0.000	0.000	0.000
Y1		0.000	0.000	0.000	0.000	0.000
Y2		0.000	0.000	0.000	0.000	0.000
Y3		0.000	0.000	0.000	0.000	0.000
IU		0.000	0.000	0.000	0.000	0.000
SU		0.855	0.098	8.716	1.027	1.027
IY		0.232	0.059	3.901	0.435	0.435
SY		0.240	0.031	7.830	1.025	1.025
Thresholds						
U0\$1		2.655	0.206	12.877		
U1\$1		2.655	0.206	12.877		
U2\$1		2.655	0.206	12.877		
U3\$1		2.655	0.206	12.877		

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

Residual Variances

Y0	0.175	0.032	5.470	0.175	0.380
Y1	0.266	0.029	9.159	0.266	0.509
Y2	0.238	0.027	8.810	0.238	0.457
Y3	0.269	0.054	5.014	0.269	0.392
IU	7.982	1.086	7.351	0.990	0.990
SU	0.685	0.202	3.400	0.988	0.988
IY	0.279	0.040	7.019	0.981	0.981
SY	0.054	0.017	3.224	0.979	0.979

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

R-Square

Observed Variable	R-Square
U0	0.710
U1	0.682
U2	0.650
U3	0.666
Y0	0.620
Y1	0.491
Y2	0.543
Y3	0.608

Latent Variable	R-Square
-----------------	----------

IU	0.010
SU	0.012
IY	0.019
SY	0.021

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

Technical 4 Output

ESTIMATED MEANS FOR THE LATENT VARIABLES					
	IU	SU	IY	SY	MALE
1	0.305	0.758	0.312	0.204	0.536

ESTIMATED COVARIANCE MATRIX FOR THE LATENT VARIABLES					
	IU	SU	IY	SY	MALE
IU	8.062				
SU	-1.170	0.694			
IY	1.214	-0.007	0.285		
SY	-0.010	0.003	-0.042	0.055	
MALE	0.142	-0.045	0.037	-0.017	0.249

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Output Excerpts Step 2 Of A Two-Part Growth Model (Continued)

ESTIMATED CORRELATION MATRIX FOR THE LATENT VARIABLES					
	IU	SU	IY	SY	MALE
IU	1.000				
SU	-0.495	1.000			
IY	0.801	-0.015	1.000		
SY	-0.014	0.016	-0.336	1.000	
MALE	0.100	-0.109	0.139	-0.145	1.000

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General Latent Variable Modeling Using Mplus Version 3

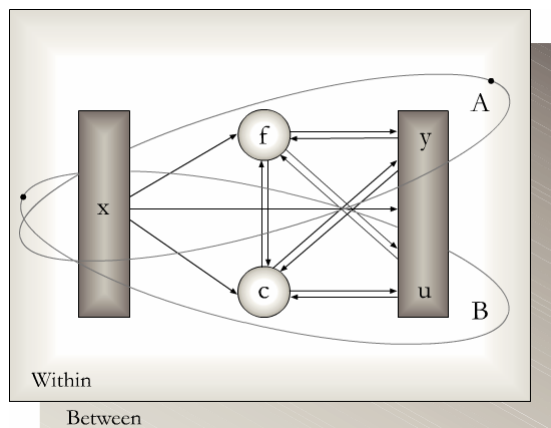
Block 3: Mixture Modeling

Bengt Muthén
bmuthen@ucla.edu

Mplus: www.statmodel.com

1

General Latent Variable Modeling Framework



2

Categorical Latent Variables

- Mixture regression
- Latent class analysis
- Latent transition analysis
- Missing data modeling

Categorical and Continuous Latent Variables

- SEMM
- Growth mixture modeling

3

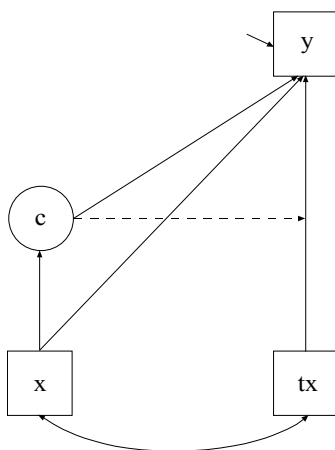
Randomized Preventive Interventions and Complier-Average Causal Effect Estimation (CACE)

- Angrist, Imbens & Rubin (1996)
- Yau & Little (1998, 2001)
- Jo (2002)
- Dunn et al. (2003)

- Compliance status observed for those invited for treatment
- Compliance status unobserved for controls

4

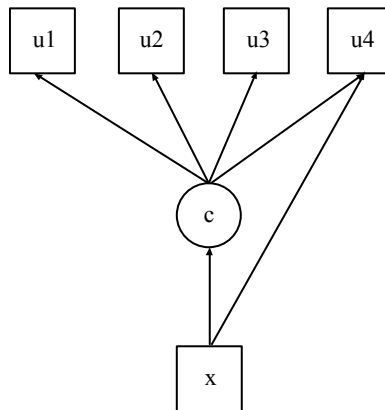
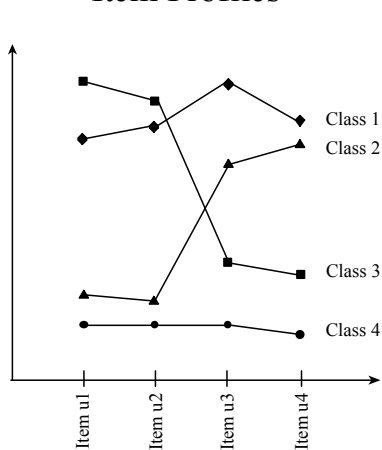
CACE Mixture Modeling



5

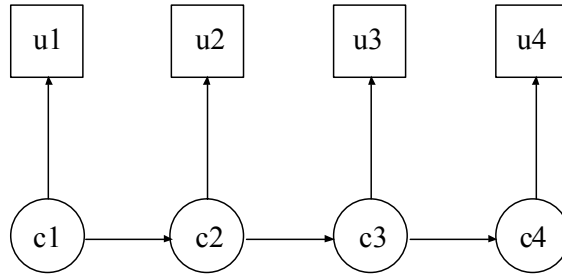
Latent Class Analysis

Item Profiles



6

Hidden Markov Modeling



Latent Transition Analysis

Transition Probabilities

Mover Class (c=1)

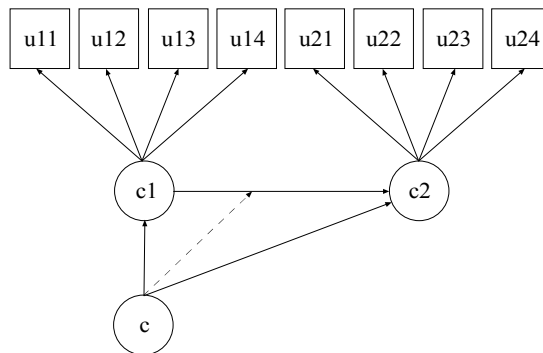
	c2	
	1	2
1	0.6	0.4
2	0.3	0.7

Stayer Class (c=2)

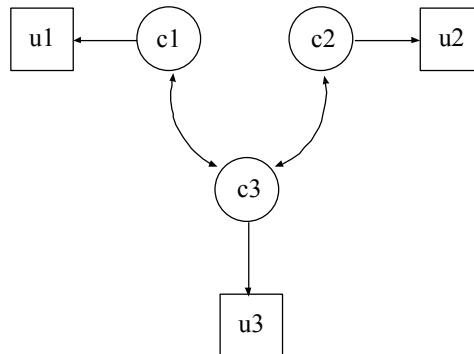
	c2	
	1	2
1	0.90	0.10
2	0.05	0.95

Time Point 1

Time Point 2



Loglinear Modeling of Frequency Tables



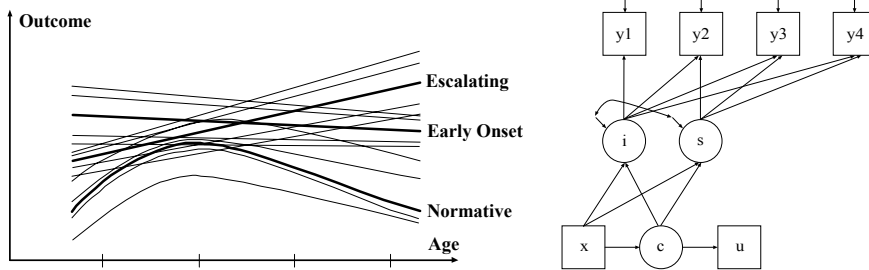
9

Growth Mixture Modeling

- Muthén, B. & Shedden, K. (1999). Finite mixture modeling with mixture outcomes using the EM algorithm. **Biometrics**, 55, 463-469.
- Muthén, B., Brown, C.H., Masyn, K., Jo, B., Khoo, S.T., Yang, C.C., Wang, C.P., Kellam, S., Carlin, J., & Liao, J. (2002). General growth mixture modeling for randomized preventive interventions. **Biostatistics**, 3, 459-475.

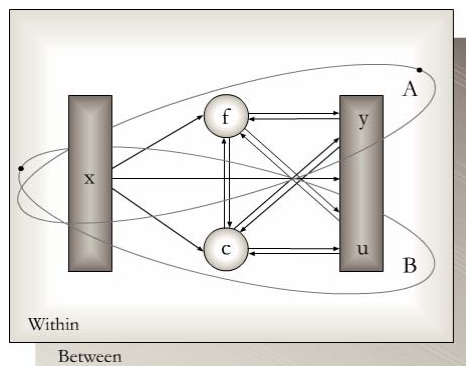
10

Growth Mixture Modeling



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General Modeling Framework



- x background variables
- y, u continuous, censored, count, and categorical outcome variables
- f continuous latent variables
- c categorical latent variables

12

Summary Of Techniques Using Latent Classes

	Outcome/ Indicator Scale	Number of Timepoints	Number of Outcomes/ Timepoint	Within-Class Variation	
				Standard	Mplus
LCA	u	Single	Multiple	No	Yes
LPA	y	Single	Multiple	No	Yes
LCFA	u, y	Single	Multiple	No	Yes
LCGA	u, y	Multiple	Single Multiple	No	Yes (GMM)
LTA	u, y	Multiple	Multiple	No	Yes

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Summary Of Techniques Using Latent Classes (Continued)

	Outcome/ Indicator Scale	Number of Timepoints	Number of Outcomes/ Timepoint	Within-Class Variation	
				Standard	Mplus
SEMM	u, y	Single	Multiple	Yes	Yes
GMM	u, y	Multiple	Single Multiple	Yes	Yes
GGMM	u, y	Multiple	Single Multiple	Yes	Yes
DTSMA	u	Multiple	Single Multiple	No	Yes
LLLCA	u, y	Single Multiple	Single Multiple	NA	Yes

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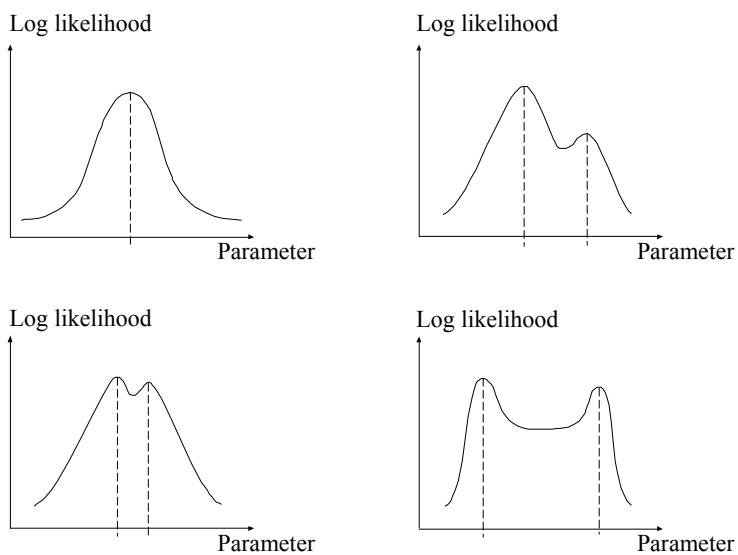
Summary Of Techniques Using Latent Classes (Continued)

LCA – Latent Class Analysis
LPA – Latent Profile Analysis
LCFA – Latent Class Factor Analysis
LCGA – Latent Class Growth Analysis
LTA – Latent Transition Analysis
SEMM – Structural Equation Mixture Modeling
GMM – Growth Mixture Modeling
GGMM – General Growth Mixture Modeling
DTSMA – Discrete-Time Survival Mixture Analysis
LLCA – Loglinear Latent Class Analysis

u – categorical dependent variables
y – continuous, censored, count dependent variables

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Global and Local Solutions



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Random Starts in Version 3

When TYPE=MIXTURE is used, random sets of starting values are generated as the default for all parameters in the model except variances and covariances. These random sets of starting values are random perturbations of either user-specified starting values or default starting values produced by the program. Maximum likelihood optimization is done in two stages. In the initial stage, 10 random sets of starting values are generated. An optimization is carried out for ten iterations using each of the 10 random sets of starting values. The ending values from the optimization with the highest loglikelihood are used as the starting values in the final stage of optimization which is carried out using the default optimization settings for TYPE=MIXTURE. Random starts can be turned off or done more thoroughly.

Recommendations for a more thorough investigation of multiple solutions:

```
STARTS = 100 10;  
or  
STARTS = 500 10;  
with  
STITERATIONS = 20;
```

17

Antisocial Behavior (ASB) Data

The Antisocial Behavior (ASB) data were taken from the National Longitudinal Survey of Youth (NLSY) that is sponsored by the Bureau of Labor Statistics. These data are made available to the public by Ohio State University. The data were obtained as a multistage probability sample with oversampling of blacks, Hispanics, and economically disadvantaged non-blacks and non-Hispanics.

Data for the analysis include 17 antisocial behavior items that were collected in 1980 when respondents were between the ages of 16 and 23 and the background variables of age, gender and ethnicity. The ASB items assessed the frequency of various behaviors during the past year. A sample of 7,326 respondents has complete data on the antisocial behavior items and the background variables of age, gender, and ethnicity. Following is a list of the 17 items:

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Antisocial Behavior (ASB) Data (Continued)

Damaged property	Use other drugs
Fighting	Sold marijuana
Shoplifting	Sold hard drugs
Stole < \$50	“Con” someone
Stole > \$50	Take auto
Use of force	Broken into building
Seriously threaten	Held stolen goods
Intent to injure	Gambling operation
Use marijuana	

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Input For LCA Of 17 Antisocial Behavior (ASB) Items With Random Starts

```
TITLE:      LCA of 17 ASB items
DATA:      FILE IS asb.dat;
           FORMAT IS 34x 42f2;

VARIABLE:  NAMES ARE property fight shoplift lt50 gt50 force
           threat injure pot drug soldpot solddrug con auto
           bldg goods gambling dsml1-dsm22 sex black hisp;
           USEVARIABLES ARE property-gambling;
           CLASSES = c(5);
           CATEGORICAL ARE property-gambling;

ANALYSIS:  TYPE = MIXTURE;
           STARTS = 500 10;
           STITERATIONS = 20;

OUTPUT:    TECH8 TECH10 TECH11;

SAVEDATA:  FILE IS asb.sav;
           SAVE IS CPROB;
```

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items With Random Starts

Loglikelihood values at local maxima and seeds:

```

-40808.314  195353
-40808.406  783165
-40808.406  863691
-40815.960  939709
-40815.960  303634
-40815.960  85734
-40815.960  316165
-40815.960  458181
-40815.960  502532
-40816.006  605161
    
```

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Deciding On The Number Of Classes For The ASB Items

Number of Classes	Loglikelihood	# par.	BIC	AIC	Entropy	LRT p-value for k-1
1	-48,168.475	17	96,488	96,370	NA	NA
2	-42,625.653	35	85,563	85,321	.838	.0000
3	-41,713.142	53	83,898	83,532	.743	.0000
4	-41,007.498	71	82,647	82,157	.742	.0000
5	-40,808.314	89	82,409	81,795	.741	.0000
6	-40,604.231	107	82,161	81,422	.723	.0019

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Deciding On The Number Of Classes For The ASB Items (Continued)

Four-Class Solution

FINAL CLASS COUNTS AND PROPORTIONS OF TOTAL SAMPLE SIZE BASED ON ESTIMATED POSTERIOR PROBABILITIES

Class 1	672.41667	0.09178	High
Class 2	1354.73100	0.18492	Drug
Class 3	1821.71706	0.24866	Person Offense
Class 4	3477.13527	0.47463	Normative (Pot)

Five-Class Solution

FINAL CLASS COUNTS AND PROPORTIONS OF TOTAL SAMPLE SIZE BASED ON ESTIMATED POSTERIOR PROBABILITIES

			Comparison To Four-Class Solution
Class 1	138.06985	0.01888	High
Class 2	860.41897	0.11771	Property Offense
Class 3	1257.56652	0.17151	Drug
Class 4	1909.32749	0.26219	Person Offense
Class 5	3160.61717	0.42971	Normative (Pot)

Six-Class Solution - adds a variation on Class 2 in the 5-class solution

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Input For LCA Of 17 Antisocial Behavior (ASB) Items

TITLE: LCA of 17 ASB items

DATA: FILE IS asb.dat;
FORMAT IS 34x 42f2;

VARIABLE: NAMES ARE property fight shoplift lt50 gt50 force
threat injure pot drug soldpot solddrug con auto
bldg goods gambling dsml-dsm22 sex black hisp;

USEVARIABLES ARE property-gambling;

CLASSES = c(4);

CATEGORICAL ARE property-gambling;

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Input For LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

ANALYSIS: TYPE = MIXTURE;

```

MODEL:
      %OVERALL%                !Not needed in Version 3
      %c#1%                    !Not needed in Version 3
      [property$1-gambling$1*0]; !Not needed in Version 3
      %c#2%                    !Not needed in Version 3
      [property$1-gambling$1*1]; !Not needed in Version 3
      %c#3%                    !Not needed in Version 3
      [property$1-gambling$1*2]; !Not needed in Version 3
      %c#4%                    !Not needed in Version 3
      [property$1-gambling$1*3]; !Not needed in Version 3
  
```

OUTPUT: TECH8 TECH10 TECH11;

```

SAVEDATA: FILE IS asb.sav;
          SAVE IS CPROB;
  
```

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Output Excerpts For LCA Of 17 Antisocial Behavior (ASB) Items

Tests of Model Fit

Loglikelihood

H0 Value	-41007.498
----------	------------

Information Criteria

Number of Free parameters	71
Akaike (AIC)	82156.996
Bayesian (BIC)	82646.838
Sample-Size Adjusted BIC	82421.215
(n* = (n + 2) / 24)	
Entropy	0.742

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Output Excerpts For LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Chi-Square Test of Model Fit for the Latent Class Indicator Model Part**

Pearson Chi-Square

Value	20827.381
Degrees of freedom	130834
P-Value	1.0000

Likelihood Ratio Chi-Square

Value	6426.411
Degrees of Freedom	130834
P-Value	1.0000

**Of the 131072 cells in the latent class indicator table, 166 were deleted in the calculation of chi-square due to extreme values.

Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Classification Information

FINAL CLASS COUNTS AND PROPORTIONS OF TOTAL SAMPLE SIZE

Class 1	672.41594	0.09178
Class 2	1354.72999	0.18492
Class 3	1821.73064	0.24867
Class 4	3477.12344	0.47463

CLASSIFICATION OF INDIVIDUALS BASED ON THEIR MOST LIKELY CLASS MEMBERSHIP

Class Counts and Proportions

Class 1	664	0.09064
Class 2	1237	0.16885
Class 3	1772	0.24188
Class 4	3653	0.49863

Average Class Probabilities by Class

	1	2	3	4
Class 1	0.896	0.057	0.046	0.000
Class 2	0.032	0.835	0.090	0.043
Class 3	0.021	0.072	0.803	0.104
Class 4	0.000	0.043	0.070	0.887

Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Number of Classes

TECHNICAL 10

UNIVARIATE MODEL FIT INFORMATION

Estimated Probabilities			
Variable	H1	H0	Residual
PROPERTY			
Category 1	0.815	0.815	0.000
Category 2	0.185	0.185	0.000
FIGHT			
Category 1	0.719	0.719	0.000
Category 2	0.281	0.281	0.000
SHOPLIFT			
Category 1	0.736	0.736	0.000
Category 2	0.264	0.264	0.000

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

BIVARIATE MODEL FIT INFORMATION

Estimated Probabilities				
Variable	Variable	H1	H0	Residual
PROPERTY		FIGHT		
Category 1	Category 1	0.635	0.631	0.004
Category 1	Category 2	0.180	0.184	-0.004
Category 2	Category 1	0.084	0.088	-0.004
Category 2	Category 2	0.101	0.097	0.004
PROPERTY		SHOPLIFT		
Category 1	Category 1	0.656	0.646	0.010
Category 1	Category 2	0.159	0.169	-0.010
Category 2	Category 1	0.080	0.090	-0.010
Category 2	Category 2	0.105	0.095	0.010

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

TECHNICAL 11- check that the H0 loglikelihood value is the same as the k-1 class H0 loglikelihood value to be certain a local solution has not been reached.

VUONG-LO-MENDELL-RUBIN LIKELIHOOD RATIO TEST FOR 3 (H0) VERSUS 4 CLASSES

H0 Loglikelihood Value	-41713.142
2 Times the Loglikelihood Difference	1411.288
Difference in the Number of Parameters	19
Mean	-0.960
Standard Deviation	43.222
P-Value	0.0000

LO-MENDELL-RUBIN ADJUSTED LRT TEST

Value	1402.991
P-Value	0.0000

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Class 1

Thresholds

PROPERTY\$1	-1.267	0.142	-8.911
FIGHT\$1	-1.047	0.117	-8.972
SHOPLIFT\$1	-1.491	0.125	-11.927
LT50\$1	-0.839	0.114	-7.377
GT50\$1	0.523	0.117	4.477
FORCE\$1	1.027	0.113	9.113
THREAT\$1	-1.495	0.125	-11.996
INJURE\$1	0.394	0.096	4.125
POT1	-2.220	0.193	-11.496
DRUG\$1	-0.394	0.122	-3.234
SOLDPOT\$1	-0.053	0.116	-0.455
SOLDDRUG\$1	1.784	0.135	13.233
CONS\$1	-0.585	0.109	-5.388
AUTO\$1	0.591	0.102	5.796
BLDG\$1	0.290	0.112	2.591
GOODS\$1	-0.697	0.112	-5.699
GAMBLING\$1	1.722	0.125	13.774

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Class 2

Thresholds

PROPERTY\$1	1.533	0.113	13.550
FIGHT\$1	1.403	0.118	11.857
SHOPLIFT\$1	0.310	0.083	3.755
LT50\$1	0.988	0.085	11.561
GT50\$1	3.543	0.218	16.252
FORCE\$1	4.058	0.319	12.718
THREAT\$1	0.499	0.097	5.153
INJURE\$1	2.462	0.165	14.881
POT1	-3.232	0.311	-10.403
DRUG\$1	-0.336	0.118	-2.853
SOLDPOT\$1	1.033	0.109	9.457
SOLDDRUG\$1	3.189	0.180	17.691
CONS\$1	1.386	0.093	14.918
AUTO\$1	2.473	0.144	17.195
BLDG\$1	3.381	0.223	15.186
GOODS\$1	2.167	0.148	14.632
GAMBLING\$1	4.078	0.269	15.158

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Class 3

Thresholds

PROPERTY\$1	0.962	0.104	9.267
FIGHT\$1	-0.134	0.089	-1.508
SHOPLIFT\$1	0.780	0.096	8.084
LT50\$1	1.350	0.108	12.470
GT50\$1	3.360	0.197	17.067
FORCE\$1	2.456	0.116	21.213
THREAT\$1	-0.747	0.105	-7.131
INJURE\$1	1.465	0.102	14.420
POT1	0.567	0.088	6.467
DRUG\$1	3.649	0.298	12.258
SOLDPOT\$1	5.393	0.737	7.320
SOLDDRUG\$1	6.263	0.752	8.325
CONS\$1	0.508	0.079	6.467
AUTO\$1	2.121	0.102	20.809
BLDG\$1	3.100	0.193	16.099
GOODS\$1	1.969	0.130	15.122
GAMBLING\$1	3.514	0.182	19.260

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Class 4

Thresholds

PROPERTY\$1	3.687	0.176	20.891
FIGHT\$1	2.281	0.107	21.345
SHOPLIFT\$1	2.609	0.114	22.923
LT50\$1	3.046	0.119	25.566
GT50\$1	5.796	0.403	14.386
FORCE\$1	5.276	0.343	15.395
THREAT\$1	2.171	0.136	15.985
INJURE\$1	5.765	0.664	8.682
POT1	1.290	0.065	19.888
DRUG\$1	4.430	0.305	14.502
SOLDPOT\$1	6.367	0.589	10.801
SOLDDRUG\$1	6.499	0.573	11.342
CONS\$1	2.525	0.106	23.928
AUTO\$1	4.314	0.208	20.784
BLDG\$1	6.741	0.739	9.120
GOODS\$1	5.880	0.611	9.627
GAMBLING\$1	6.816	0.954	7.144

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

LATENT CLASS INDICATOR MODEL PART IN PROBABILITY SCALE

Class 3

PROPERTY			
Category 2	0.277	0.021	13.321
FIGHT			
Category 2	0.533	0.022	24.193
SHOPLIFT			
Category 2	0.314	0.021	15.120
LT50			
Category 2	0.206	0.018	11.635
GT50			
Category 2	0.034	0.006	5.256
FORCE			
Category 2	0.079	0.008	9.379
THREAT			
Category 2	0.678	0.023	29.703
INJURE			
Category 2	0.188	0.015	12.118

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

POT				
Category 2	0.362	0.020	17.887	
DRUG				
Category 2	0.025	0.007	3.447	
SOLDPOT				
Category 2	0.005	0.003	1.364	
SOLDDRUG				
Category 2	0.002	0.001	1.332	
CON				
Category 2	0.376	0.018	20.388	
AUTO				
Category 2	0.107	0.010	10.989	
BLDG				
Category 2	0.043	0.008	5.427	
GOODS				
Category 2	0.122	0.014	8.752	
GAMBLING				
Category 2	0.029	0.005	5.645	

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

Technical 8 Output

E STEP	ITER	LOGLIKELIHOOD	ABS CHANGE	REL CHANGE	CLASS COUNTS
1	-0.50814249D+05	0.0000000	0.0000000		888.234 1659.562 2208.576 2569.628
2	-0.41810482D+05	9003.7666995	0.1771898		831.905 1722.366 2165.174 2606.555
3	-0.41706620D+05	103.8616123	0.0024841		767.588 1807.168 2146.235 2605.009
4	-0.41657122D+05	49.4986699	0.0011868		714.379 1867.660 2146.792 2597.170
5	-0.41623995D+05	33.1269450	0.0007952		671.621 1905.257 2162.382 2586.740

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

96	-0.41007499D+05	0.0002095	0.0000000	673.025	1354.398
				1824.820	3473.758
97	-0.41007499D+05	0.0001814	0.0000000	672.982	1354.419
				1824.606	3473.993
98	-0.41007499D+05	0.0001572	0.0000000	672.943	1354.439
				1824.408	3474.211
99	-0.41007499D+05	0.0001362	0.0000000	672.906	1354.457
				1824.222	3474.414
100	-0.41007499D+05	0.0001180	0.0000000	672.872	1354.475
				1824.050	3474.604

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Output Excerpts LCA Of 17 Antisocial Behavior (ASB) Items (Continued)

153	-0.41007498D+05	0.0000001	0.0000000	672.424	1354.725
				1821.771	3477.081
154	-0.41007498D+05	0.0000001	0.0000000	672.423	1354.726
				1821.767	3477.085
155	-0.41007498D+05	0.0000000	0.0000000	672.422	1354.726
				1821.764	3477.088
171	-0.41007498D+05	0.0000000	0.0000000	672.416	1354.730
				1821.733	3477.121
172	-0.41007498D+05	0.0000000	0.0000000	672.416	1354.730
				1821.732	3477.122
173	-0.41007498D+05	0.0000000	0.0000000	672.416	1354.730
				1821.731	3477.123

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Class Probability Excerpts LCA Of 17 Antisocial Behavior (ASB) Items

Saved Data And Posterior Class Probabilities

```

0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0.
.000      .001      .013      .987      4.000
1. 0. 0. 1. 0. 0. 0. 1. 1. 1. 0. 0. 0. 0. 0.
.005      .995      .000      .000      2.000
0. 1. 0. 0. 0. 0. 1. 0. 0. 0. 0. 1. 1. 0. 1. 0.
.003      .001      .996      .000      3.000
0. 0. 0. 0. 0. 0. 1. 0. 0. 0. 0. 0. 0. 0. 0. 0.
.000      .004      .191      .805      4.000
0. 1. 0. 0. 0. 0. 1. 0. 1. 0. 0. 0. 0. 1. 0. 0.
.004      .121      .871      .004      3.000
    
```

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Class Table 1. LCA and EFA for Antisocial Behavior (n=7326)

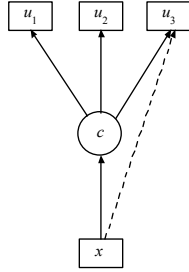
	LCA Solution				EFA Solution		
	Categorical Factors				Continuous Factors		
	C#1	C#2	C#3	C#4	Factor 1	Factor 2	Factor 3
Property	0.78	0.18	0.28	0.02	0.65	0.19	-0.04
Fighting	0.74	0.20	0.53	0.09	0.19	0.60	-0.13
Shoplifting	0.82	0.42	0.31	0.07	0.61	-0.03	0.18
Stole < \$50	0.70	0.27	0.21	0.05	0.85	-0.21	0.05
Stole > \$50	0.37	0.03	0.03	0.00	0.81	0.00	0.01
Use of force	0.26	0.02	0.08	0.01	0.34	0.37	-0.01
Seriously threaten	0.82	0.38	0.68	0.10	-0.11	0.89	0.03
Intent to injure	0.40	0.08	0.19	0.00	-0.11	0.83	0.08
Use marijuana	0.90	0.96	0.36	0.22	-0.02	0.00	0.88
Use other drugs	0.60	0.58	0.03	0.01	0.01	-0.02	0.88
Sold marijuana	0.51	0.26	0.01	0.00	0.15	0.07	0.74
Sold hard drugs	0.14	0.04	0.00	0.00	0.19	0.09	0.59
“Con” someone	0.64	0.20	0.38	0.07	0.43	0.25	-0.07
Take auto	0.36	0.08	0.11	0.01	0.45	0.15	0.07
Broken into bldg.	0.43	0.03	0.04	0.00	0.80	0.03	0.01
Held stolen goods	0.67	0.10	0.12	0.00	0.69	0.11	0.06
Gambling operation	0.15	0.02	0.03	0.00	0.28	0.36	0.08
Class Prob.	0.09	0.18	0.25	0.47			

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LCA with Covariates

Dichotomous indicators $u: u_1, u_2, \dots, u_r$. Categorical latent variable $c: c = k; k = 1, 2, \dots, K$. Marginal probability for item $u_j = 1$,

$$P(u_j = 1) = \sum_{k=1}^K P(c = k) P(u_j = 1 | c = k). \quad (5)$$



With a covariate x , consider $P(u_j = 1 | c = k, x), P(c = k | x)$,

$$\text{logit} [P(u_j = 1 | c = k, x)] = \lambda_{jk} + \kappa_j x, \quad (6)$$

$$\text{logit} [P(c = k | x)] = \alpha_k + \gamma_k x. \quad (7)$$

Multinomial Logistic Regression Of c On x

The multinomial logistic regression model expresses the probability that individual i falls in class k of the latent class variable c as a function of the covariate x ,

$$P(c_i = k | x_i) = \frac{e^{\alpha_k + \gamma_k x_i}}{\sum_{s=1}^K e^{\alpha_s + \gamma_s x_i}}, \quad (87)$$

where $\alpha_K = 0, \gamma_K = 0$ so that $e^{\alpha_K + \gamma_K x_i} = 1$.

This implies that the log odds comparing class k to the last class K is

$$\log[P(c_i = k | x) / P(c_i = K | x)] = \alpha_k + \gamma_k x_i. \quad (88)$$

Input For LCA Of 9 Antisocial Behavior (ASB Items With Covariates)

```
TITLE:      LCA of 9 ASB items with three covariates

DATA:      FILE IS asb.dat;
           FORMAT IS 34x 51f2;

VARIABLE:  NAMES ARE property fight shoplift lt50 gt50 force
           threat injure pot drug soldpot solddrug con auto
           bldg goods gambling dsml-dsm22 male black hisp
           single
           divorce dropout college onset f1 f2 f3 age94;
           USEVARIABLES ARE property fight shoplift lt50
           threat pot drug con goods age94 male black;
           CLASSES = c(4);
           CATEGORICAL ARE property-goods;

ANALYSIS:  TYPE = MIXTURE;
```

45

Input For LCA Of 9 Antisocial Behavior (ASB Items With Covariates) (Continued)

```
MODEL:

%OVERALL%

c#1-c#3 ON age94 male black;

%c#1%
[property$1-gambling$1*0]; !Not needed in Version 3
!Not needed in Version 3

%c#2%
[property$1-gambling$1*1]; !Not needed in Version 3
!Not needed in Version 3

%c#3%
[property$1-gambling$1*2]; !Not needed in Version 3
!Not needed in Version 3

%c#4%
[property$1-gambling$1*3]; !Not needed in Version 3
!Not needed in Version 3

OUTPUT:   TECH8;
```

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Output Excerpts LCA Of 9 Antisocial Behavior (ASB) Items With Covariates

Tests of Model Fit

Loglikelihood

H0 Value	-30416.942
Information Criteria	
Number of Free parameters	48
Akaike (AIC)	60929.884
Bayesian (BIC)	61261.045
Sample-Size Adjusted BIC	61108.512
(n* = (n + 2) / 24)	
Entropy	0.690

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Output Excerpts LCA Of 9 Antisocial Behavior (ASB) Items With Covariates (Continued)

LATENT CLASS REGRESSION MODEL PART

C#1	ON			
AGE94		-.285	.028	-10.045
MALE		2.578	.151	17.086
BLACK		.158	.139	1.141
C#2	ON			
AGE94		.069	.022	3.182
MALE		.187	.110	1.702
BLACK		-.606	.139	-4.357
C#3	ON			
AGE94		-.317	.028	-11.311
MALE		1.459	.101	14.431
BLACK		.999	.117	8.513
Intercepts				
C#1		-1.822	.174	-10.485
C#2		-.748	.103	-7.258
C#3		-.324	.125	-2.600

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Output Excerpts LCA Of 9 Antisocial Behavior (ASB) Items With Covariates (Continued)

Classification Information

FINAL CLASS COUNTS AND PROPORTIONS OF TOTAL SAMPLE SIZE

Class 1	928.40043	0.12673
Class 2	1499.08913	0.20463
Class 3	2249.50562	0.30706
Class 4	2649.00482	0.36159

CLASSIFICATION OF INDIVIDUALS BASED ON THEIR MOST LIKELY CLASS MEMBERSHIP

Class Counts and Proportions

Class 1	920	0.12558
Class 2	1433	0.19560
Class 3	2154	0.29402
Class 4	2819	0.38479

Average Class Probabilities by Class

	1	2	3	4
Class 1	0.859	0.065	0.076	0.000
Class 2	0.047	0.808	0.087	0.058
Class 3	0.033	0.067	0.816	0.084
Class 4	0.000	0.048	0.105	0.846

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Output Excerpts LCA Of 9 Antisocial Behavior (ASB) Items With Covariates (Continued)

LATENT CLASS INDICATOR MODEL PART

Class 1

PROPERTY\$1	-1.185	0.133	-8.883
FIGHT\$1	-0.967	0.108	-8.958
SHOPLIFT\$1	-1.307	0.130	-10.078
LT50\$1	-0.647	0.108	-5.998
THREAT\$1	-1.383	0.108	-12.847
POT\$1	-1.656	0.151	-10.935
DRUG\$1	0.093	0.106	0.878
CON\$1	-0.384	0.091	-4.201
GOODS\$1	-0.299	0.106	-2.816

Class 2

PROPERTY\$1	1.834	0.126	14.583
FIGHT\$1	1.700	0.144	11.769
SHOPLIFT\$1	0.425	0.079	5.418
LT50\$1	1.113	0.082	13.504
THREAT\$1	0.549	0.095	5.782
POT\$1	-2.561	0.246	-10.415
DRUG\$1	-0.127	0.107	-1.185
CON\$1	1.346	0.092	14.595
GOODS\$1	2.272	0.137	16.592

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Output Excerpts LCA Of 9 Antisocial Behavior (ASB) Items With Covariates (Continued)

LATENT CLASS INDICATOR MODEL PART

Class 3

PROPERTY\$1	1.494	0.110	13.632
FIGHT\$1	0.032	0.085	0.378
SHOPLIFT\$1	1.312	0.100	13.083
LT50\$1	1.929	0.122	15.858
THREAT\$1	-0.184	0.090	-2.049
POT\$1	0.916	0.097	9.484
DRUG\$1	4.484	0.480	9.345
CON\$1	0.937	0.074	12.626
GOODS\$1	2.668	0.152	17.502

Class 4

PROPERTY\$1	4.699	0.433	10.840
FIGHT\$1	3.988	0.512	7.787
SHOPLIFT\$1	2.943	0.153	19.222
LT50\$1	3.192	0.158	20.193
THREAT\$1	2.929	0.219	13.360
POT\$1	1.443	0.082	17.624
DRUG\$1	5.236	0.637	8.225
CON\$1	2.814	0.145	19.402
GOODS\$1	7.307	1.901	3.844

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Calculating Latent Class Probabilities For Different Covariate Values

For a class,

$$\text{logit} = \text{intercept} + b1*\text{age94} + b2*\text{male} + b3*\text{black}$$

Example 1: For age94 = 0, male = 0, black = 0

where age94 = 0 is age 16

male = 1 is female

black = 0 is not black

	exp	probability (exp/sum)
logitc1 = -1.822	0.162	0.069
logitc2 = -0.748	0.473	0.201
logitc3 = -0.324	0.723	0.307
logitc4 = 0	1.0	0.424
sum	2.358	1.001

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Calculating Latent Class Probabilities For Different Covariate Values (Continued)

Example 2: For age94 = 1, male = 1, black = 1
 where age94 = 1 is age 17
 male = 1 is male
 black = 1 is black

$$\begin{aligned} \text{logitc1} &= -1.822 + (-0.285*1) + (2.578*1) + (0.158*1) \\ &= 0.629 \end{aligned}$$

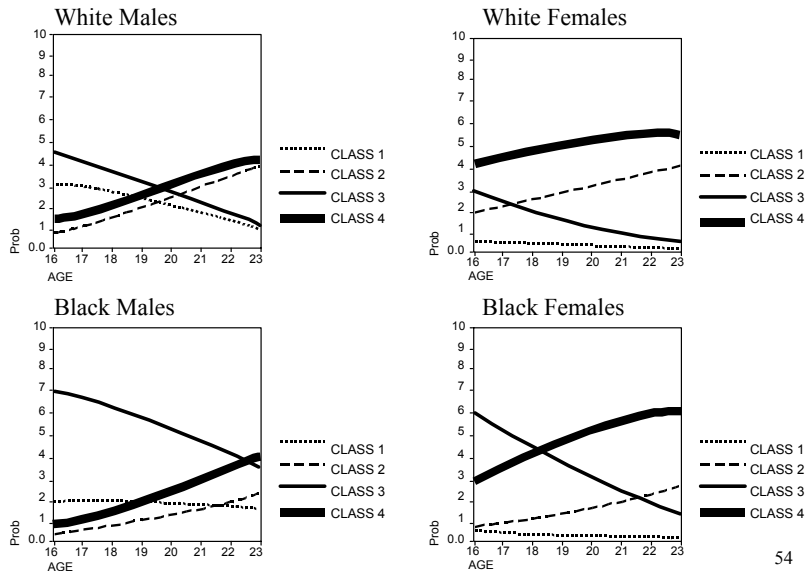
$$\begin{aligned} \text{logitc2} &= -0.748 + 0.069*1 + 0.187*1 + (-0.606*1) \\ &= -1.098 \end{aligned}$$

$$\begin{aligned} \text{logitc3} &= -0.324 + (-0.317*1) + 1.459*1 + 0.999*1 \\ &= 1.817 \end{aligned}$$

	exp	probability (exp/sum)
logitc1 = 0.629	1.876	0.200
logitc2 = -1.098	0.334	0.036
logitc3 = 1.817	6.153	0.657
logitc4 = 0	1.0	0.107
sum	9.363	1.000

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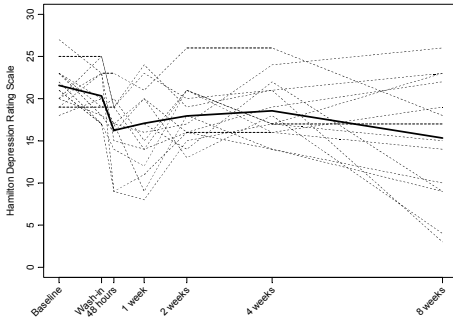
ASB Classes Regressed on Age, Male, Black in the NLSY (n=7326)



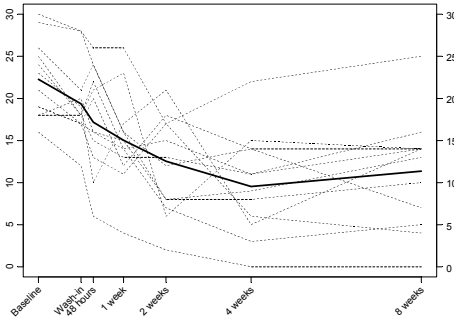
54

A Clinical Trial of Depression Medication: 2-Class Growth Mixture Modeling

Placebo Non-Responders, 55%



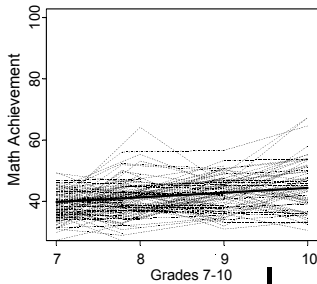
Placebo Responders, 45%



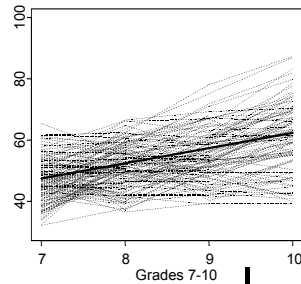
55

Growth Mixture Modeling: LSAY Math Achievement Trajectory Classes and the Prediction of High School Dropout

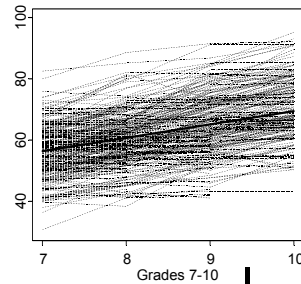
Poor Development: 20%



Moderate Development: 28%



Good Development: 52%



Dropout:	69%	8%	
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General Latent Variable Modeling Using Mplus Version 3

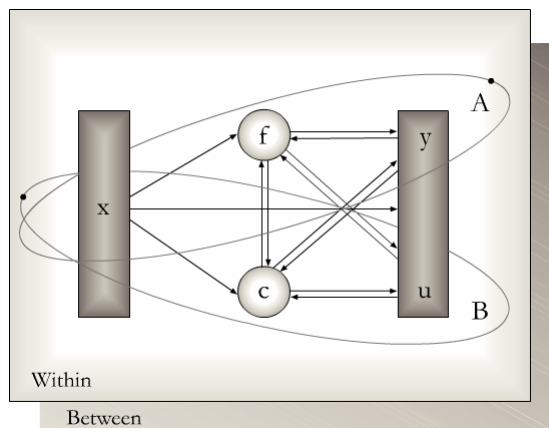
Block 4: Multilevel Modeling

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Mplus: www.statmodel.com

1

General Latent Variable Modeling Framework



2

Multilevel Modeling with Continuous and Categorical Latent Variables

- Multilevel regression
- Multilevel CFA, SEM
- Multilevel growth modeling
- Multilevel discrete-time survival analysis

- Multilevel regression mixture analysis (CACE)
- Multilevel latent class analysis
- Multilevel growth mixture modeling

3

Multilevel Regression Analysis With Random Intercepts And Random Slopes In Multilevel Terms

Two-level analysis (individual i in cluster j):

y_{ij} : individual-level outcome variable

x_{ij} : individual-level covariate

w_j : cluster-level covariate

Random intercepts, random slopes:

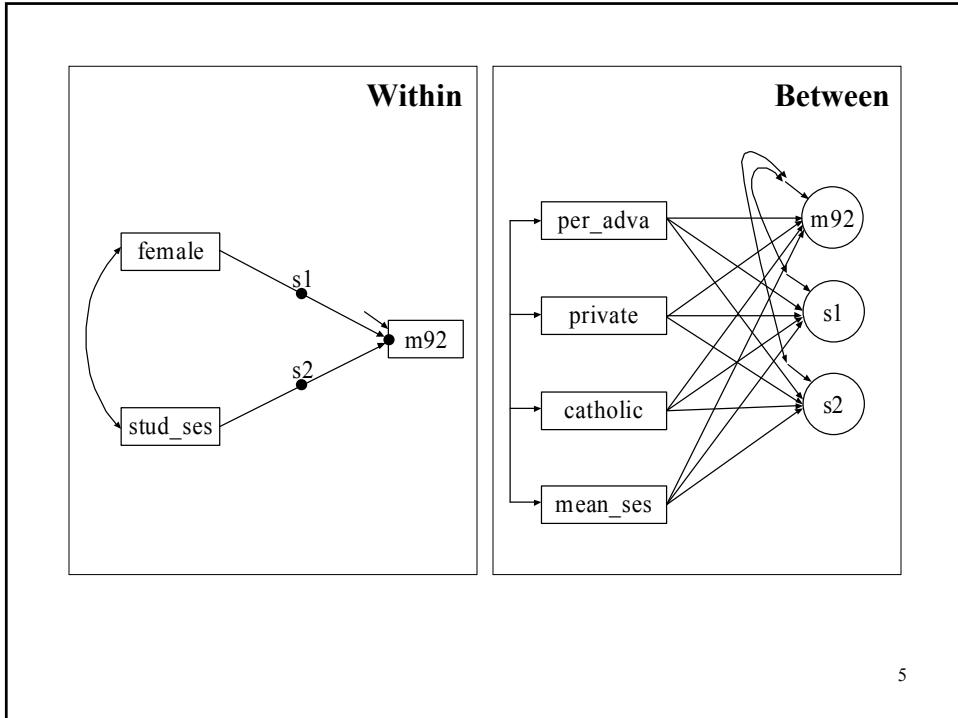
$$\text{Level 1 (Within)} : y_{ij} = \beta_{0j} + \beta_{1j} x_{ij} + r_{ij}, \quad (8)$$

$$\text{Level 2 (Between)} : \beta_{0j} = \gamma_{00} + \gamma_{01} w_j + u_{0j}, \quad (9)$$

$$\text{Level 2 (Between)} : \beta_{1j} = \gamma_{10} + \gamma_{11} w_j + u_{1j}, \quad (10)$$

- Mplus gives the same estimates as HLM/MLwiN ML (not REML): $V(r)$ (residual variance for level 1), $\gamma_{00}, \gamma_{01}, \gamma_{10}, \gamma_{11}, V(u_0), V(u_1), Cov(u_0, u_1)$
- Centering of x : subtracting grand mean or group (cluster) mean
- Model testing with varying covariance structure, marginal covariance matrix for y

4



5

Input For Multilevel Regression Model

```

TITLE:      multilevel regression

DATA:      FILE IS completev2.dat;
           ! National Education Longitudinal Study (NELS)
           FORMAT IS f8.0 12f5.2 f6.3 f11.4 23f8.2
           f18.2 f8.0 4f8.2;

VARIABLE:  NAMES ARE school r88 m88 s88 h88 r90 m90 s90 h90 r92
           m92 s92 h92 stud ses f2pnlwt transfer minor coll_asp
           algebra retain aCa back female per_min0 hw_time
           salary dis_fair clās_dis mean_col per_high_unsafe
           num_frie teaqual par_invo ac_track urban size rural
           private mean_ses catholic stu_teach per_adva tea_exce
           tea_res;

           USEV = m92 female stud_ses per_adva private catholic
           mean_ses;

           !per_adva = percent teachers with an MA or higher

WITHIN = female stud_ses;
BETWEEN = per_adva private catholic mean_ses;
MISSING = school;
CLUSTER = school;
CENTERING = GRANDMEAN (stud_ses);

```

6

Input For Multilevel Regression Model

```
ANALYSIS: TYPE = TWOLEVEL RANDOM MISSING;

MODEL:
    %WITHIN%
    s1 | m92 ON female;
    s2 | m92 ON stud_ses;

    %BETWEEN%
    s1 WITH m92; s2 WITH m92;
    m92 s1 s2 ON per_adva private catholic mean_ses;

OUTPUT:  TECH8 SAMPSTAT;
```

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Output Excerpts For Multilevel Regression Model

N = 10,933

Summary of Data

Number of clusters 902

Size (s) Cluster ID with Size s

1	89863	75862	52654	1995	32661	89239	56214	
2	41743	81263	45025	26790	60281	82860	56241	21474
	4570	27159	11662	87842	38454			
3	65407	61407	83048	42640	41412	67708	83085	39685
	40402	93469	98582	68595	11517	17543	75498	81069
	66512							
4	31646	68153	85508	26234	83390	60835	7400	20770
	5095	10904	93569	38063	86733	66125	51670	10910
	98461	44395	95317	64112	50880	77381	12835	47555
	9208	93859	35719	67574	20048	34139	25784	80675
5	14464	74791	18219	10468	72193	97616	15773	877
	9471	83234	68254	68028	70718	3496	6842	45854

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Output Excerpts For Multilevel Regression Model

S2	ON			
PER_ADVA		1.348	0.521	2.587
PRIVATE		-1.890	0.706	-2.677
CATHOLIC		-1.467	0.562	-2.612
MEAN_SES		1.031	0.283	3.640
M92	ON			
PER_ADVA		0.195	0.727	0.268
PRIVATE		1.505	1.108	1.358
CATHOLIC		0.765	0.650	1.178
MEAN_SES		3.912	0.399	9.814
S1	WITH			
M92		-4.456	1.007	-4.427
S2	WITH			
M92		0.128	0.399	0.322
Intercepts				
M92		54.886	0.428	128.231
S1		-0.856	0.507	-1.688
S2		4.075	0.309	13.208
Residual Variances				
M92		8.679	1.003	8.649
S1		5.740	1.411	4.066
S2		0.307	0.527	0.583

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Random Slopes

- In single-level modeling random slopes β_i describe variation across individuals i ,

$$y_i = \alpha_i + \beta_i x_i + \varepsilon_i, \quad (100)$$

$$\alpha_i = \alpha + \zeta_{0i}, \quad (101)$$

$$\beta_i = \beta + \zeta_{1i}, \quad (102)$$

Resulting in heteroscedastic residual variances

$$V(y_i | x_i) = V(\beta_i) x_i^2 + \theta. \quad (103)$$

- In two-level modeling random slopes β_j describe variation across clusters j

$$y_{ij} = \alpha_j + \beta_j x_{ij} + \varepsilon_{ij}, \quad (104)$$

$$\alpha_j = \alpha + \zeta_{0j}, \quad (105)$$

$$\beta_j = \beta + \zeta_{1j}, \quad (106)$$

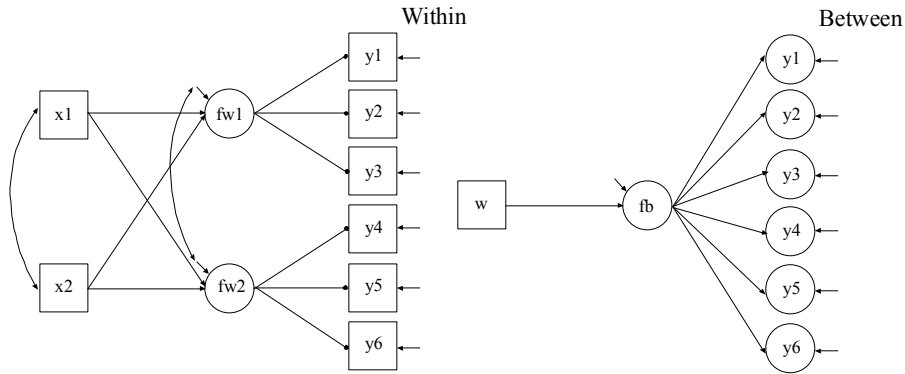
A small variance for a random slope typically leads to slow convergence of the ML-EM iterations. This suggests respecifying the slope as fixed.

Mplus allows random slopes for predictors that are

- Observed covariates
- Observed dependent variables (Version 3)
- Continuous latent variables (Version 3)

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Two-Level Factor Analysis with Covariates



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Three-Level Modeling in Multilevel Terms

Time point t , individual i , cluster j .

- y_{ij} : individual-level, outcome variable
- a_{1ij} : individual-level, time-related variable (age, grade)
- a_{2ij} : individual-level, time-varying covariate
- x_{ij} : individual-level, time-invariant covariate
- w_j : cluster-level covariate

Three-level analysis (Mplus considers Within and Between)

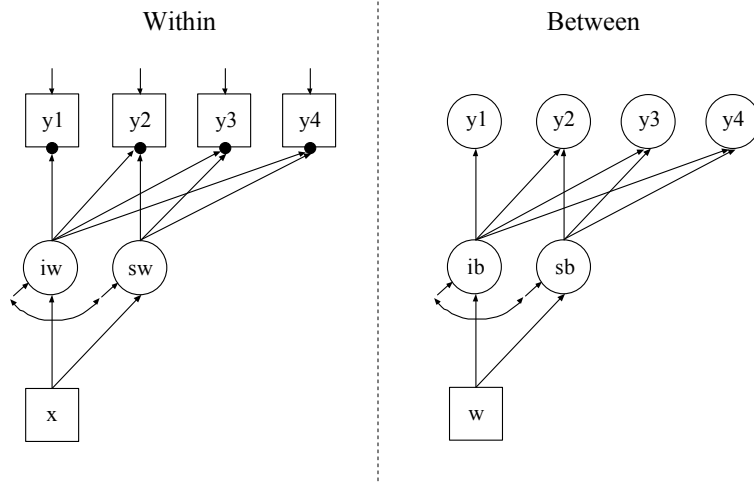
$$\text{Level 1 (Within)} : y_{ij} = \pi_{0ij} + \pi_{1ij} a_{1ij} + \pi_{2ij} a_{2ij} + e_{ij}, \quad (1)$$

$$\text{Level 2 (Within)} : \begin{cases} \pi_{0ij} = \beta_{00j} + \beta_{01j} x_{ij} + r_{0ij}, \\ \pi_{1ij} = \beta_{10j} + \beta_{11j} x_{ij} + r_{1ij}, \\ \pi_{2ij} = \beta_{20j} + \beta_{21j} x_{ij} + r_{2ij}. \end{cases} \quad (2)$$

$$\text{Level 3 (Between)} : \begin{cases} \beta_{00j} = \gamma_{000} + \gamma_{001} w_j + u_{00j}, \\ \beta_{10j} = \gamma_{100} + \gamma_{101} w_j + u_{10j}, \\ \beta_{20j} = \gamma_{200t} + \gamma_{201t} w_j + u_{20j}, \\ \beta_{01j} = \gamma_{010} + \gamma_{011} w_j + u_{01j}, \\ \beta_{11j} = \gamma_{110} + \gamma_{111} w_j + u_{11j}, \\ \beta_{21j} = \gamma_{210} + \gamma_{211} w_j + u_{21j}. \end{cases} \quad (3)$$

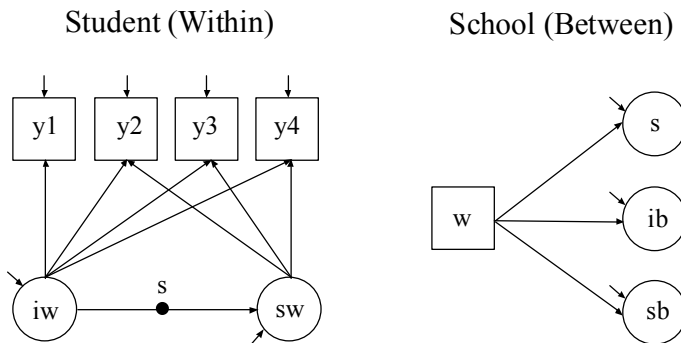
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Two-Level Growth Modeling (3-Level Modeling)



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Multilevel Modeling with a Random Slope for Latent Variables



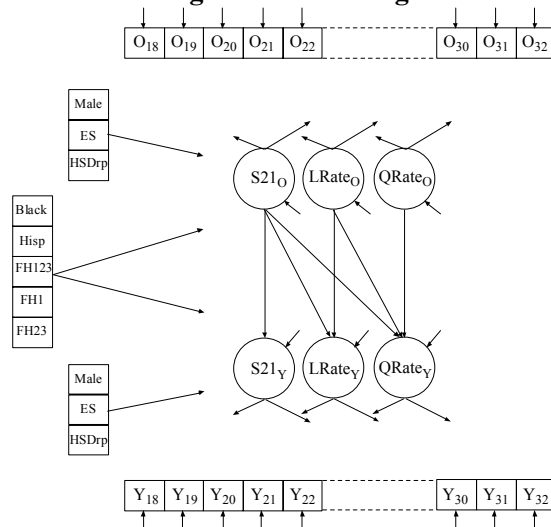
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Multivariate Modeling of Family Members

- Multilevel modeling: clusters independent, model for between- and within-cluster variation, units within a cluster statistically equivalent
- Multivariate approach: clusters independent, model for all variables for each cluster unit, different parameters for different cluster units.
 - used in the latent variable growth modeling, where the cluster units are the repeated measures over time
 - allows for different cluster sizes by missing data techniques
 - more flexible than the multilevel approach, but computationally convenient only for applications with small cluster sizes (e.g. twins, spouses)

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Figure 1. A Longitudinal Growth Model of Heavy Drinking for Two-Sibling Families

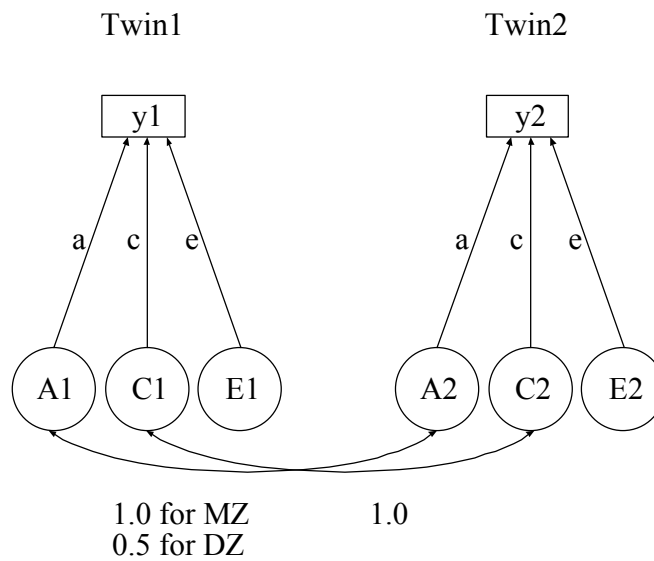


Source: Khoo, S.T. & Muthen, B. (2000). Longitudinal data on families: Growth modeling alternatives. *Multivariate Applications in Substance Use Research*, J. Rose, L. Chassin, C. Presson & J. Sherman (eds.), Hillsdale, N.J.: Erlbaum, pp. 43-78.

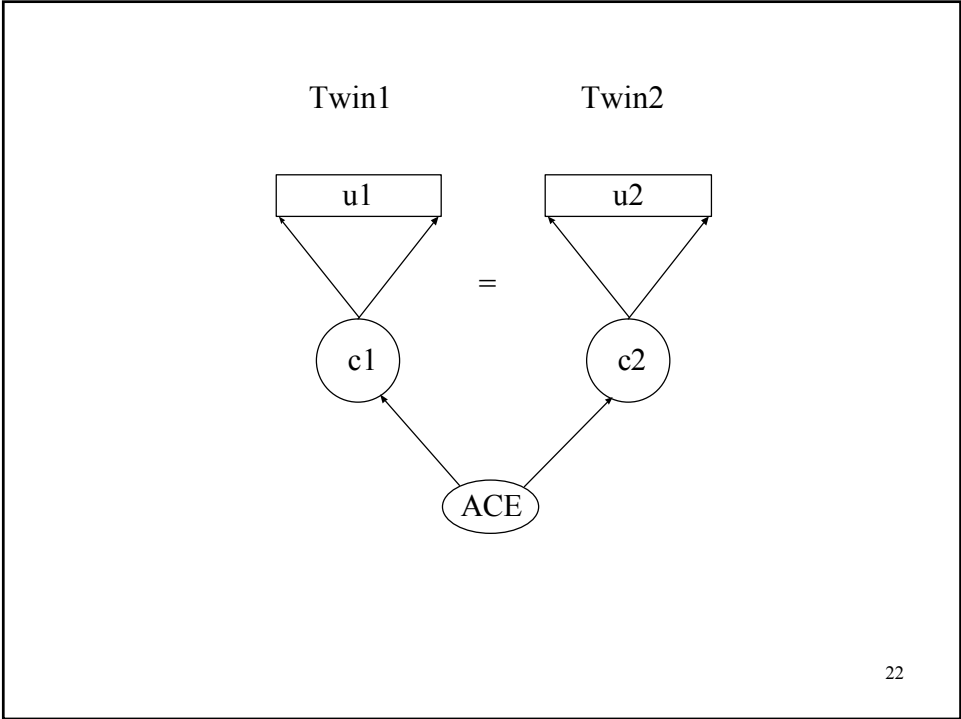
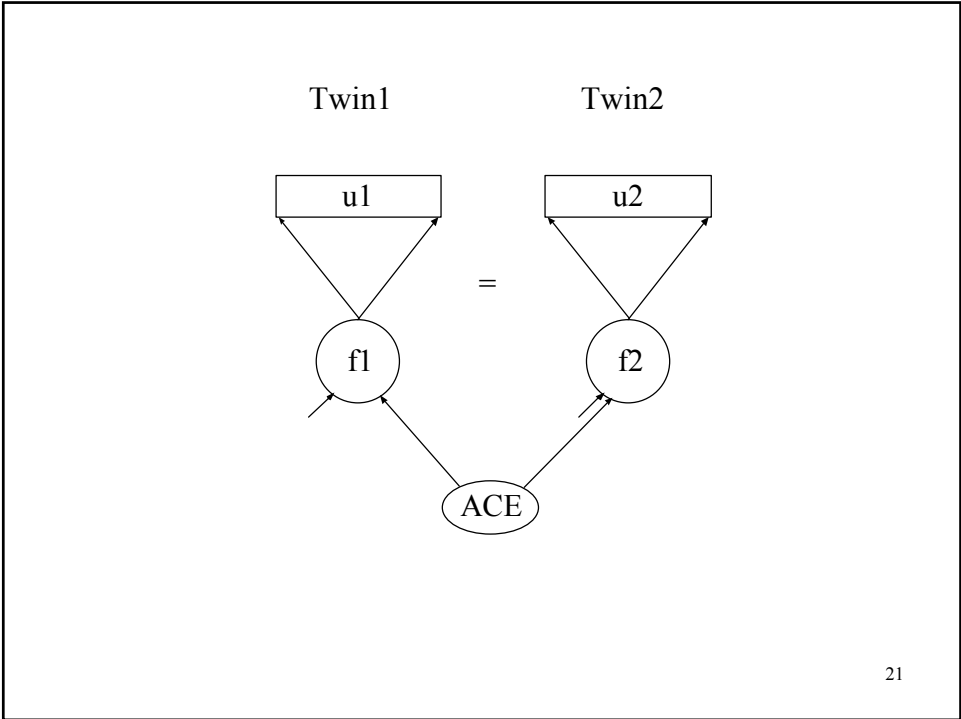
18

Twin Modeling

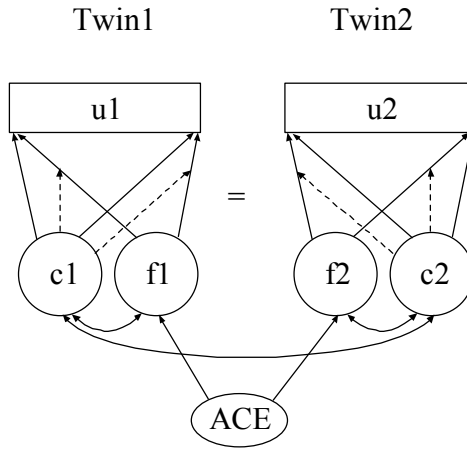
19



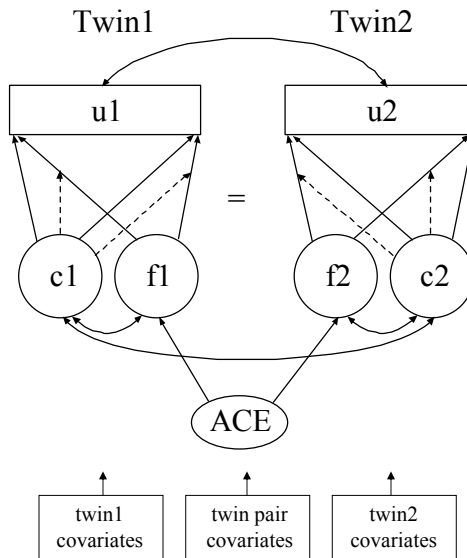
20



Hybrid Model (Severity LCA or Three-Part Modeling)



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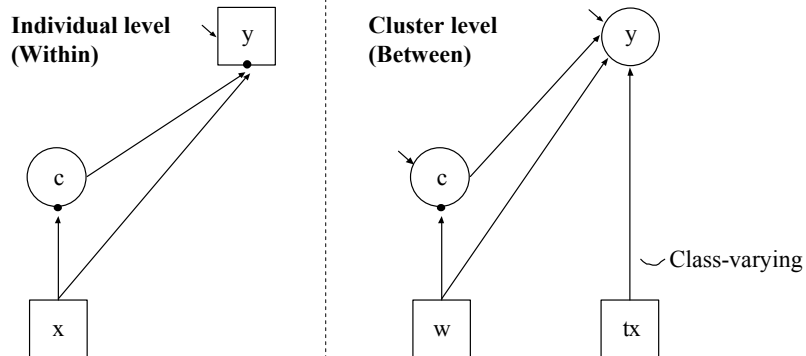


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Multilevel Mixture Modeling

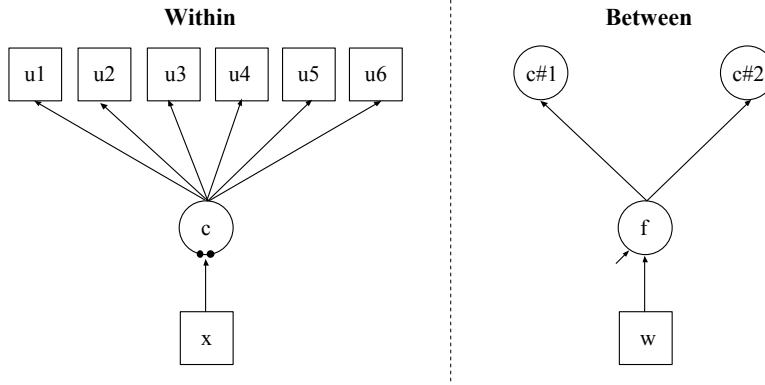
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Two-Level CACE Mixture Modeling



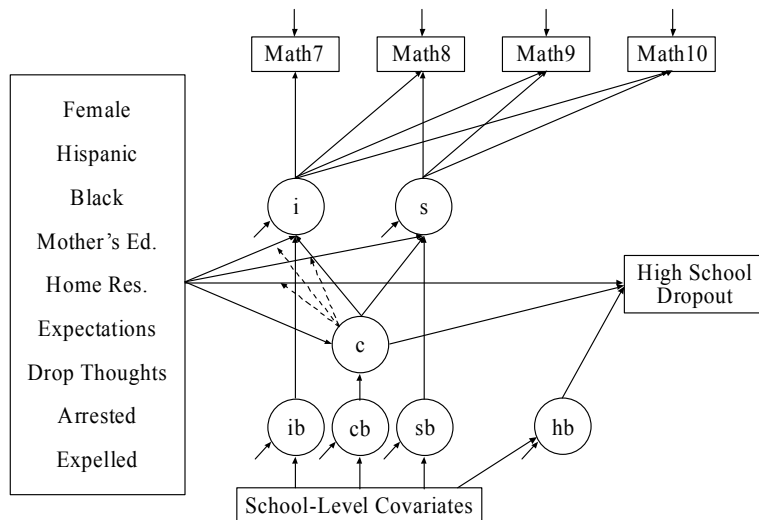
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Two-Level Latent Class Analysis



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Multilevel Growth Mixture Modeling



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Monte Carlo Simulations in Mplus

- Data generation, analysis, and results summaries across replications
- Studies of tests of model fit, parameter estimation, standard errors, coverage, and power as a function of model variations, parameter values, and sample size
- Model Population, Model Missing, Model for analysis
- Full modeling framework available: continuous and categorical latent variables, multilevel data, different types of outcomes

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References

- See the Mplus web site www.statmodel.com

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General Latent Variable Modeling Framework

